

Applied Time Series Econometrics

Johannes W. Fedderke
ERSA and University of Cape Town

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Establishing the Univariate Characteristics of the Data

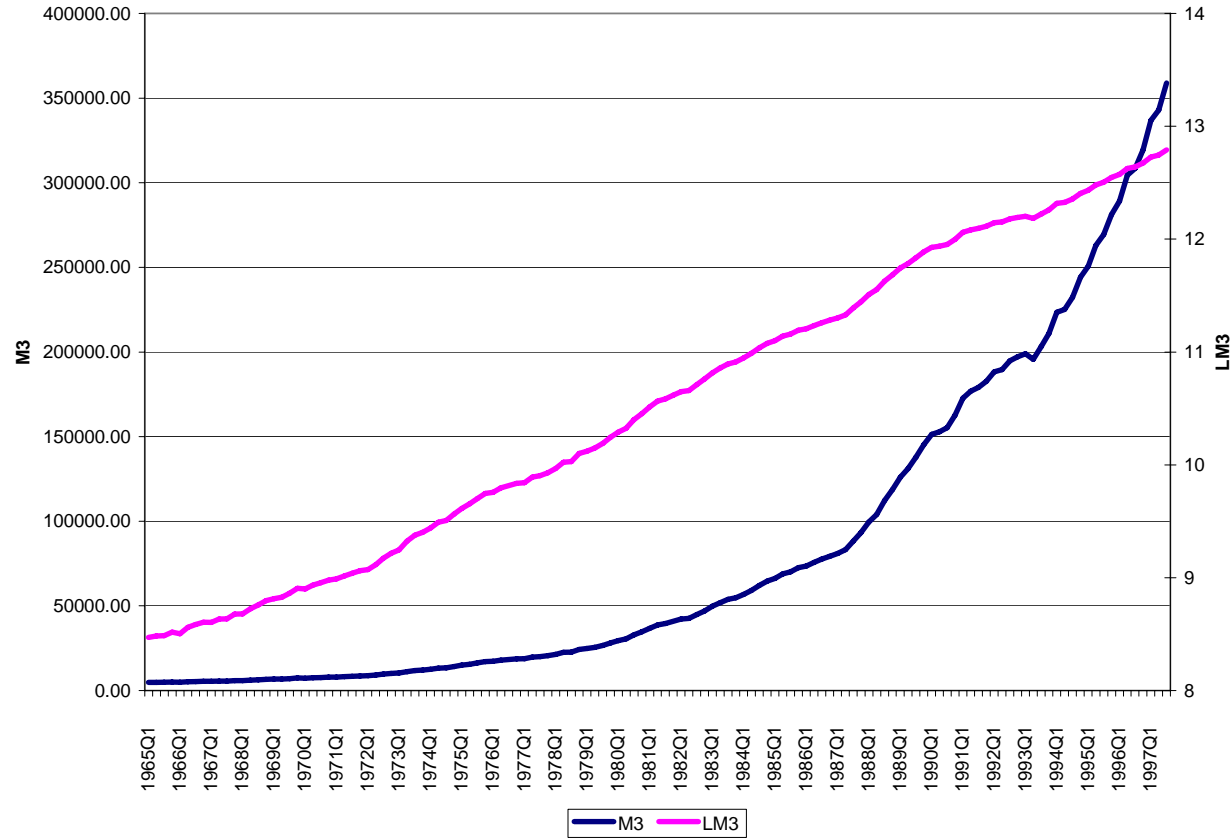
Have demonstrated the importance of stationarity of data to estimation.

Now:

- *Formal* tests which examine the stationarity properties of the data to be used.
- Rest on an examination of residuals of DGP's.
- Data transformations.

1. Establishing the Appropriate Variable Transformations

Consider:



- Note: data series *should* be integrated in only one of the two scales.
- Illustration: macroeconomic series often grow over time subject to some stochastic shocks. Hence:

$$\begin{aligned}
 Y_t &= (1 + \gamma) Y_{t-1} e_t, \\
 e_t &\sim ILN \left(1, \sigma^2 \right) \\
 \gamma &= \text{constant}
 \end{aligned}$$

constant.

- It follows:

$$\begin{aligned}
 \ln Y_t &= \ln (1 + \gamma) + \ln Y_{t-1} + \ln e_t \\
 \text{or } y_t &\approx \gamma + y_{t-1} + \epsilon_t \quad \epsilon_t \sim IN \left(0, \sigma^2 \right) \\
 \therefore \Delta y_t &= \gamma + \epsilon_t
 \end{aligned}$$

\implies variable grows at a constant proportional rate of γ , subject to random shocks.

\implies For (most) macroeconomic series growth rate varies about a (typically constant) mean,

$\implies y_t \sim I(1), \Delta y_t \sim I(0)$.

- But then it is not possible for $Y_t \sim I(1)$, since:

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + u_t, \quad \alpha_1 = 1$$
$$\implies \Delta Y_t = \alpha_0 + u_t$$

if $\Delta Y_t \sim I(0)$,

$$\implies \Delta Y_t \sim (\alpha_0, \sigma^2)$$

\implies *declining* proportional growth rate in the series over time.

- It is not possible for both $\Delta y_t \sim I(0)$ and $\Delta Y_t \sim I(0)$ to be true.

- Are there formal tests available?
- Yes:
 - Ermini-Hendry
 - Schwarz Bayesian.

1.1 Ermini-Hendry Test

- Banerjee, Dolado, Galbraith and Hendry 1993:192f
- Estimate:

$$\Delta y_t = \mu + \varepsilon_t$$

to formulate:

$$\lambda = \mu + (\sigma^2/2)$$

- Then estimate:

$$\Delta Y_t = \gamma + \sum \beta_i \Delta Y_{t-i} + \delta \exp(\lambda t)$$

- Lag length chosen to eliminate serial correlation.
- Test: log-transform encompasses model in levels if:

$$\gamma = 0, \delta \neq 0$$

- Intuition:

- if the data are generated by the logarithmic model:

$$\Delta y_t = \mu + \varepsilon_t$$

- model in levels would be characterized by drift:

$$\delta \exp(\lambda t)$$

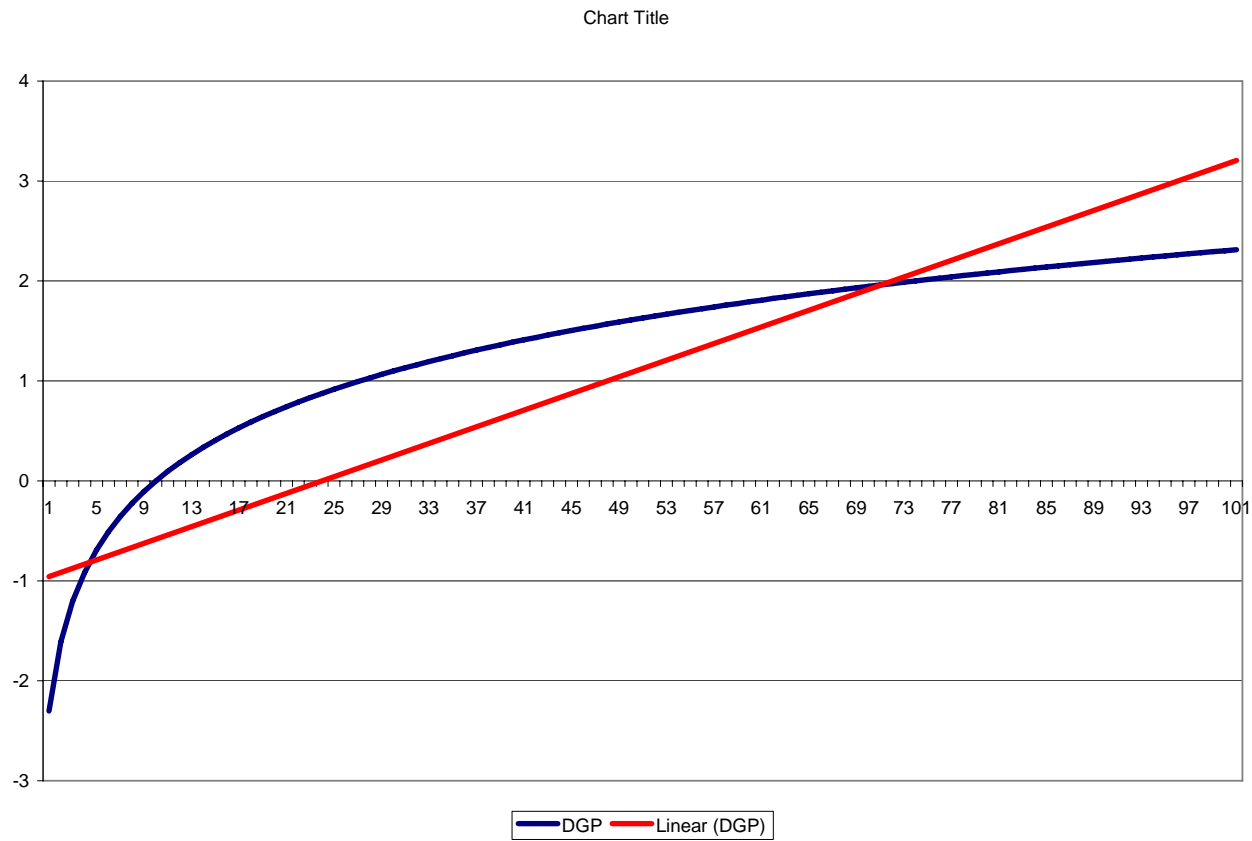
and variance:

$$\phi \exp[(2\lambda + \sigma^2)t]$$

exponentially increasing with time.

\implies drift in model with normal scale *solely* due to underlying log-DGP.

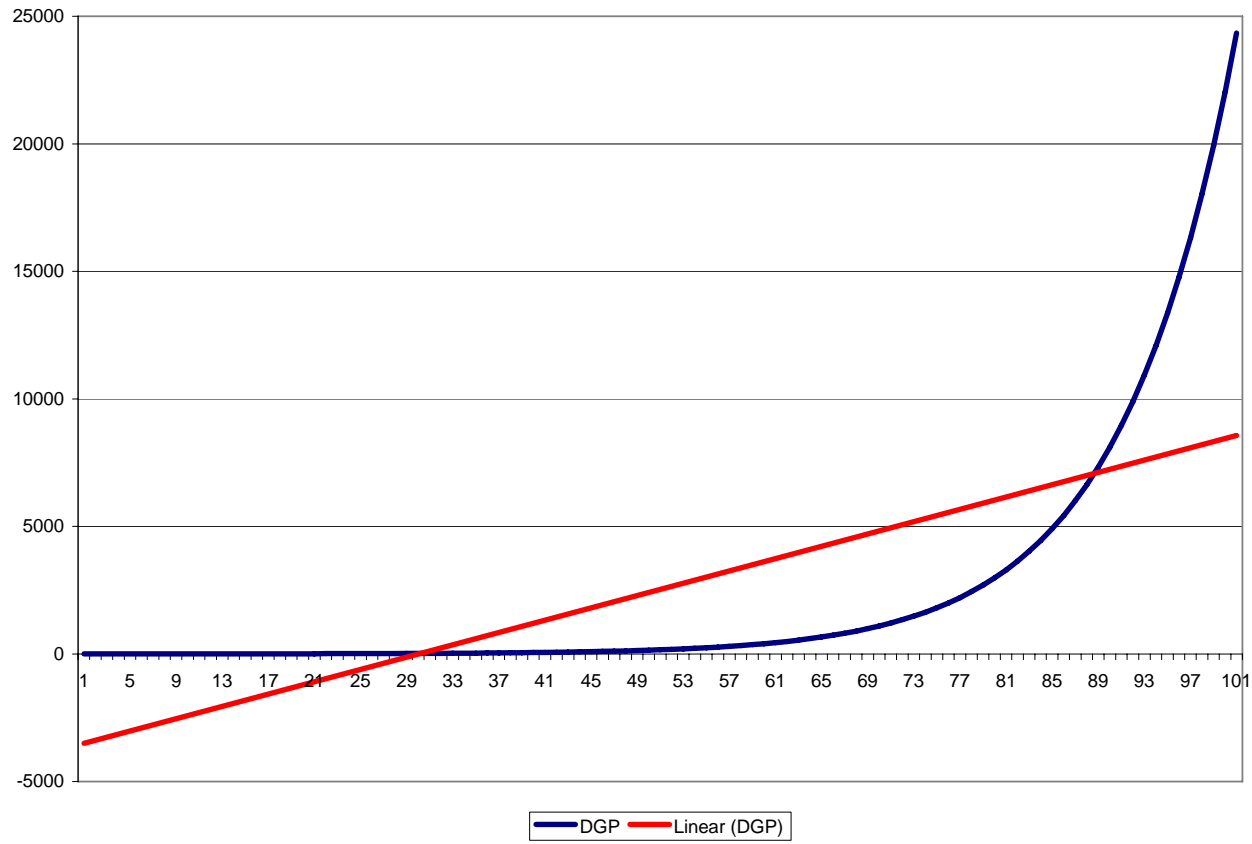
- For instance:



$$\implies t \rightarrow \infty, \sigma^2 \rightarrow \infty$$

$$\implies t \rightarrow \infty, \gamma < 0, \text{ negative drift}$$

● Similarly:



$$\implies t \rightarrow \infty, \sigma^2 \rightarrow \infty$$

$$\implies t \rightarrow \infty, \gamma > 0, \text{ positive drift}$$

- Example: South African M3
- Estimate:

$$\ln \left(\frac{M3}{P} \right)_t = \mu + \epsilon_t$$

● Gives:

```

                          Ordinary Least Squares Estimation
*****
Dependent variable is DLM3
130 observations used for estimation from 1965Q2 to 1997Q3
*****
Regressor          Coefficient          Standard Error          T-Ratio[Prob]
C                  .033239                .0016539                20.0968[.000]
*****
R-Squared          .0000          R-Bar-Squared          .0000
S.E. of Regression .018858          F-stat.                *NONE*
Mean of Dependent Variable .033239          S.D. of Dependent Variable .018858
Residual Sum of Squares .045874          Equation Log-likelihood 332.2484
Akaike Info. Criterion 331.2484          Schwarz Bayesian Criterion 329.8146
DW-statistic       1.7340
*****

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                          Diagnostic Tests
*****
*   Test Statistics   *   LM Version   *   F Version   *
*****
*   A:Serial Correlation*CHSQ( 4)= 16.7788[.002]*F( 4, 125)= 4.6311[.002]*
*   *   *   *   *   *   *
*   B:Functional Form *CHSQ( 1)= *NONE* *F( 1, 128)= *NONE*
*   *   *   *   *   *   *
*   C:Normality      *CHSQ( 2)= .77190[.680]* Not applicable
*   *   *   *   *   *   *
*   D:Heteroscedasticity*CHSQ( 1)= *NONE* *F( 1, 128)= *NONE*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

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- With:

$$\begin{aligned}\mu &= 0.033239 \\ \hat{\sigma}^2 &= (0.018858)^2\end{aligned}$$

- Hence:

$$\lambda = 0.033239 + \frac{(0.018858)^2}{2}$$

- And then estimate:

$$\Delta M3_t = \gamma + \sum \beta_i \Delta M3_{t-i} + \delta \exp(\lambda t) + \epsilon_t$$

● Gives:

```

Ordinary Least Squares Estimation
*****
Dependent variable is DM3
117 observations used for estimation from 1968Q3 to 1997Q3
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                   72.6649          244.6070            .29707[.767]
LAMBT              117.2140         28.6640             4.0892[.000]
DM3(-1)            .081582          .096559             .84489[.400]
DM3(-2)            .11356           .096418             1.1778[.242]
DM3(-3)            -.073332         .10371              -.70709[.481]
DM3(-4)            .40219           .10064              3.9964[.000]
DM3(-5)            .9469E-3         .11639              .0081354[.994]
DM3(-6)            .076589          .11832              .64730[.519]
DM3(-7)            .029385          .11986              .24516[.807]
DM3(-8)            -.49709          .12006              -4.1404[.000]
DM3(-9)            .42534           .13104              3.2459[.002]
DM3(-10)           -.46963          .11832              -3.9692[.000]
DM3(-11)           -.017022         .12360              -.13772[.891]
DM3(-12)           .41276           .12591              3.2782[.001]
DM3(-13)           -.36764          .12332              -2.9812[.004]
*****
R-Squared           .78501           R-Bar-Squared       .75550
S.E. of Regression  1860.8           F-stat.             F( 14, 102) 26.6028[.000]
Mean of Dependent Variable 3014.0           S.D. of Dependent Variable 3763.3
Residual Sum of Squares 3.53E+08           Equation Log-likelihood -1038.9
Akaike Info. Criterion -1053.9           Schwarz Bayesian Criterion -1074.6
DW-statistic        1.9701
*****

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Diagnostic Tests
*****
* Test Statistics * LM Version * F Version *
*****
* A:Serial Correlation*CHSQ( 4)= 13.5114[.009]*F( 4, 98)= 3.1987[.016]*
*
* B:Functional Form *CHSQ( 1)= .22310[.637]*F( 1, 101)= .19296[.661]*
*
* C:Normality *CHSQ( 2)= 72.2382[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 10.6655[.001]*F( 1, 115)= 11.5346[.001]*
*****

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A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

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● And:

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                          Ordinary Least Squares Estimation
*****
Dependent variable is DM3
117 observations used for estimation from 1968Q3 to 1997Q3
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                  62.4133          239.6178            .26047[.795]
LAMBT             111.1418         25.8468             4.3000[.000]
DM3(-1)           .081521          .080442             1.0134[.313]
DM3(-2)           .13806          .083829             1.6469[.103]
DM3(-4)           .40574          .094189             4.3078[.000]
DM3(-8)           -.49268         .10911              -4.5156[.000]
DM3(-9)           .40444          .10345              3.9095[.000]
DM3(-10)          -.43492         .098781            -4.4029[.000]
DM3(-12)          .40715          .11309              3.6002[.000]
DM3(-13)          -.33830         .10606              -3.1898[.002]
*****
R-Squared          .78244          R-Bar-Squared       .76414
S.E. of Regression 1827.7          F-stat.             F( 9, 107) 42.7571[.000]
Mean of Dependent Variable 3014.0          S.D. of Dependent Variable 3763.3
Residual Sum of Squares 3.57E+08          Equation Log-likelihood -1039.6
Akaike Info. Criterion -1049.6          Schwarz Bayesian Criterion -1063.4
DW-statistic       1.9648
*****

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```

                          Diagnostic Tests
*****
*   Test Statistics   *   LM Version   *   F Version   *
*****
*   A:Serial Correlation*CHSQ( 4)= 5.2854[.259]*F( 4, 103)= 1.2183[.308]*
*   *   *   *   *   *   *
*   B:Functional Form *CHSQ( 1)= .0028727[.957]*F( 1, 106)= .0026027[.959]*
*   *   *   *   *   *   *
*   C:Normality      *CHSQ( 2)= 80.4360[.000]*   Not applicable   *
*   *   *   *   *   *   *
*   D:Heteroscedasticity*CHSQ( 1)= 8.6122[.003]*F( 1, 115)= 9.1376[.003]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

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- Thus:

$$\gamma = \frac{62.4133}{(239.6178)} = 0$$

$$\delta = \frac{111.1418}{(25.8468)} \neq 0$$

- Hence: ln-transform verified.

1.2 SC Criterion

- Estimate:

$$\Delta Y_t = \gamma + \sum \beta_i \Delta Y_{t-i} + u_t$$

$$\Delta y_t = \mu + \sum \theta_i \Delta y_{t-i} + \varepsilon_t$$

- To obtain Schwarz-Bayesian criterion.
- Choose the transformation with the *larger SC* criterion.

- Consider:

$$\Delta \left(\frac{M3}{P} \right)_t = \gamma + \sum \beta_i \Delta \left(\frac{M3}{P} \right)_{t-i} + u_t$$

$$\Delta \ln \left(\frac{M3}{P} \right)_t = \mu + \sum \theta_i \Delta \ln \left(\frac{M3}{P} \right)_{t-i} + \varepsilon_t$$

● Gives:

```

Ordinary Least Squares Estimation
*****
Dependent variable is DM3
117 observations used for estimation from 1968Q3 to 1997Q3
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                   165.3344         257.0008            .64332[.521]
DM3(-1)             .23007           .078309             2.9379[.004]
DM3(-2)             .32303           .077558             4.1650[.000]
DM3(-4)             .53164           .096499             5.5093[.000]
DM3(-8)             -.52314          .11736              -4.4575[.000]
DM3(-9)             .52514           .10733              4.8927[.000]
DM3(-10)            -.33494          .10349              -3.2365[.002]
DM3(-12)            .60875           .11093              5.4875[.000]
DM3(-13)            -.29748          .11387              -2.6125[.010]
*****
R-Squared           .74484           R-Bar-Squared       .72594
S.E. of Regression  1970.1          F-stat.             F( 8, 108)         39.4084[.000]
Mean of Dependent Variable  3014.0          S.D. of Dependent Variable  3763.3
Residual Sum of Squares  4.19E+08        Equation Log-likelihood  -1048.9
Akaike Info. Criterion  -1057.9         Schwarz Bayesian Criterion  -1070.3
DW-statistic        1.9762
*****

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Diagnostic Tests
*****
* Test Statistics *          LM Version          *          F Version          *
*****
* A:Serial Correlation*CHSQ( 4)= 6.1647[.187]*F( 4, 104)= 1.4461[.224]*
*
* B:Functional Form *CHSQ( 1)= .28564[.593]*F( 1, 107)= .26187[.610]*
*
* C:Normality *CHSQ( 2)= 31.0074[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 14.9678[.000]*F( 1, 115)= 16.8702[.000]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

● And:

```

                          Ordinary Least Squares Estimation
*****
Dependent variable is DLM3
117 observations used for estimation from 1968Q3 to 1997Q3
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                  .021382          .0068833            3.1063[.002]
DLM3(-1)          .11994          .093627             1.2810[.203]
DLM3(-2)          .14808          .093723             1.5799[.117]
DLM3(-4)          .23550          .096459             2.4415[.016]
DLM3(-8)          -.11895         .099531             -1.1951[.235]
DLM3(-9)          .039329         .093783             .41936[.676]
DLM3(-10)         -.061249        .091167             -.67182[.503]
DLM3(-12)         .10055          .093256             1.0782[.283]
DLM3(-13)         -.078852        .091112             -.86544[.389]
*****
R-Squared          .12840          R-Bar-Squared       .063840
S.E. of Regression .017487         F-stat.             F( 8, 108)          1.9888[.055]
Mean of Dependent Variable .034748         S.D. of Dependent Variable .018074
Residual Sum of Squares .033026         Equation Log-likelihood 312.0828
Akaike Info. Criterion 303.0828        Schwarz Bayesian Criterion 290.6530
DW-statistic       2.0086
*****

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```

                          Diagnostic Tests
*****
*   Test Statistics   *   LM Version   *   F Version   *
*****
*   A:Serial Correlation*CHSQ( 4)= 2.2718[.686]*F( 4, 104)= .51483[.725]*
*   *               *   *               *   *
*   B:Functional Form *CHSQ( 1)= .38415[.535]*F( 1, 107)= .35247[.554]*
*   *               *   *               *   *
*   C:Normality      *CHSQ( 2)= 1.3869[.500]*   Not applicable   *
*   *               *   *               *   *
*   D:Heteroscedasticity*CHSQ( 1)= .028659[.866]*F( 1, 115)= .028176[.867]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

- Such that:

$$\Delta \left(\frac{M3}{P} \right)_t = \gamma + \sum \beta_i \Delta \left(\frac{M3}{P} \right)_{t-i} + u_t$$

$$\implies SC = -1070.3$$

$$\Delta \ln \left(\frac{M3}{P} \right)_t = \mu + \sum \theta_i \Delta \ln \left(\frac{M3}{P} \right)_{t-i} + \varepsilon_t$$

$$\implies SC = 290.6530$$

- Hence: ln-transform verified.

2. Testing for Stationarity: Unit Root Tests

Recall the central problem is that in:

$$y_t = \rho y_{t-1} + \varepsilon_t$$

$$\rho = 1$$

\implies non-stationarity

Here: test for $\rho = 1$.

Important points:

1. DGP

- (a) may include a time trend: stochastic or deterministic.
- (b) may be more complex than a simple $AR(1)$ process; may involve MA -terms.

2. Unit root tests have problems with their “power”:

- (a) in small samples they are likely to **over**-accept the null-hypothesis of a unit root.
I.e. standard unit root tests will **under**-reject the presence of a unit root.

3. Structural breaks: “standard” unit root tests lead to an **under**-rejection of the null of the presence of a unit root.
4. In the case of higher frequency (eg. quarterly) data: check for seasonal unit roots.

Overview:

1. Use of autocorrelation function.
2. Use of spectrum.
3. Standard Tests:
 - (a) (Augmented) Dickey-Fuller.
 - (b) Phillips-Perron
 - (c) Perron
 - (d) Seasonal unit roots

2.1 Autocorrelation Function

- The l' th order autocorrelation coefficient R_l for a time series is given by:

$$R_l = \frac{C_l}{C_0}, \quad l = 1, 2, \dots, \frac{n}{3}$$

$$\text{where } C_l = n^{-1} \sum_{t=l+1}^n (x_t - \bar{x})(x_{t-l} - \bar{x})$$

and $n \equiv$ sample size

- With approximate estimate of the standard error of R_l of:

$$\widehat{se}(R_l) \approx \left(n^{-1} \left[1 + 2 \sum_{j=1}^{l-1} R_j^2 \right] \right)^{\frac{1}{2}}, \quad l = 1, 2, \dots, \frac{n}{3}$$

• Hence:

$C_0 \equiv$ variance

$C_1 \equiv$ covariance between x_2 and x_1

....

$C_l \equiv$ covariance between x_{l+1} and x_1

\implies under stationarity, autocovariance between two values depends only on distance between them

- Box-Pierce and Ljung-Box statistics:
 - test for the significance of the autocorrelation coefficients of the time series *jointly*,
- Difference:
 - Ljung-Box corrects for degrees of freedom;
 - Box-Pierce does not.
 - \implies Ljung-Box performs better in small samples

- The Box-Pierce Q -statistic of order p is given by:

$$Q = n \sum_{j=1}^p R_j^2 \stackrel{a}{\sim} \chi_p^2$$

- The Ljung-Box statistic of order p is obtained from:

$$Q^* = n(n+2) \sum_{j=1}^p \frac{R_j^2}{(n-j)} \stackrel{a}{\sim} \chi_p^2$$

- Autocorrelation coefficients:
 - direct evidence on the nature of the ρ coefficient that has been central to our discussion of non-stationarity.
 - As long as the $R_l < 1$, we have sample evidence in favour of $\rho < 1$, hence stationarity.
 - Where $R_l = 1$ statistically, we have sample evidence in favour of $\rho = 1$, and hence non-stationarity.

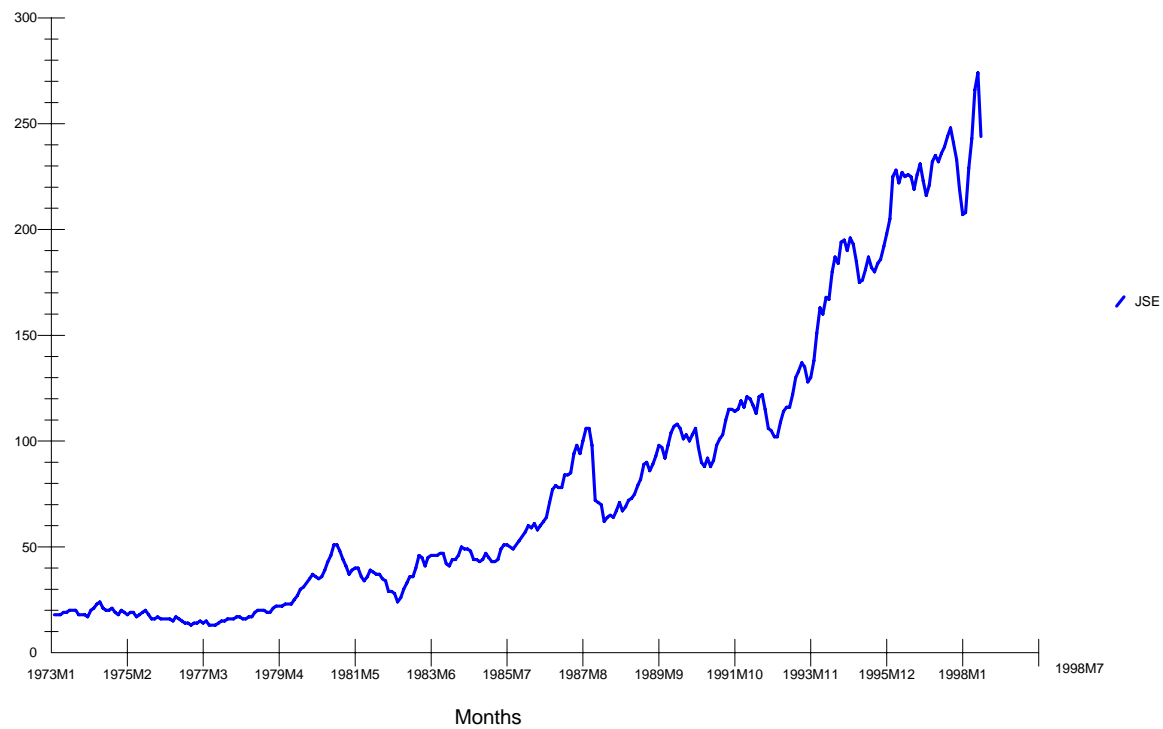
- Note:

- R_l will tend to underestimate ρ .
- Kiewitz and Phillips (1989) show that:

$$R_l \approx \rho - \frac{k(t)}{t}$$

- Bias will diminish as $t \rightarrow \infty$,
- but $k(t)$ will also increase as $t \rightarrow \infty$.
- Need: strong increase in t to eliminate the bias.
- Given the size of most time series samples, therefore, the tendency will be to underestimate ρ .
- Bootstrap: inadequate because themselves biased.

- EXAMPLE: Monthly Johannesburg Stock Exchange Index: 1973M1 1997M7



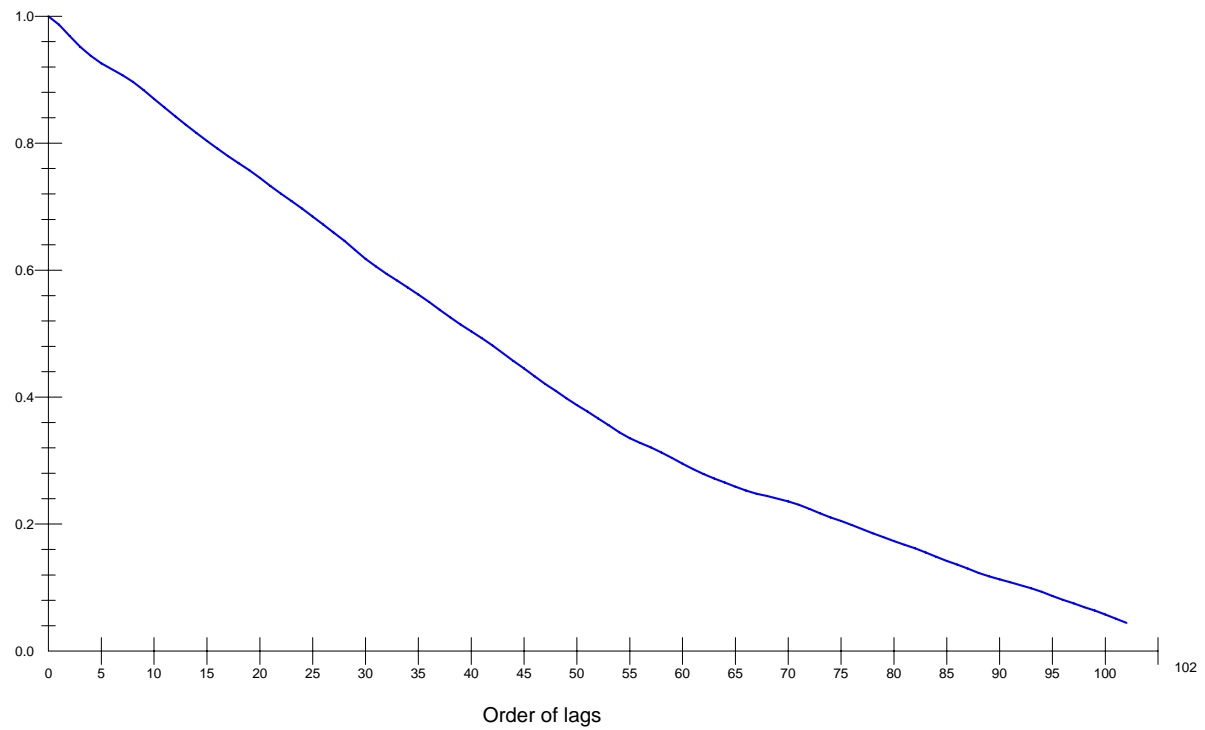
Gives:

```
Variable JSE          Sample from 1973M2 to 1998M7
*****
Order  Autocorrelation  Standard  Box-Pierce  Ljung-Box
      Coefficient      Error    Statistic   Statistic
*****
```

Order	Autocorrelation Coefficient	Standard Error	Box-Pierce Statistic	Ljung-Box Statistic
1	.98687	.057166	298.0201[.000]	300.9514[.000]
2	.96878	.098150	585.2093[.000]	591.9195[.000]
3	.95181	.12557	862.4253[.000]	873.7100[.000]
4	.93760	.14727	1131.4[.000]	1148.1[.000]
5	.92572	.16563	1393.7[.000]	1416.4[.000]
6	.91633	.18176	1650.6[.000]	1680.2[.000]
7	.90703	.19627	1902.3[.000]	1939.5[.000]
8	.89640	.20952	2148.2[.000]	2193.6[.000]
9	.88367	.22170	2387.2[.000]	2441.4[.000]
10	.87003	.23293	2618.8[.000]	2682.4[.000]
11	.85604	.24332	2843.0[.000]	2916.6[.000]
12	.84232	.25297	3060.1[.000]	3144.0[.000]
13	.82911	.26197	3270.5[.000]	3365.1[.000]
14	.81618	.27041	3474.3[.000]	3580.1[.000]
15	.80377	.27835	3672.0[.000]	3789.4[.000]
16	.79145	.28583	3863.7[.000]	3993.0[.000]
17	.77954	.29291	4049.7[.000]	4191.1[.000]
18	.76847	.29961	4230.4[.000]	4384.4[.000]
19	.75742	.30598	4405.9[.000]	4572.8[.000]
20	.74526	.31205	4575.9[.000]	4755.8[.000]
21	.73239	.31781	4740.0[.000]	4933.2[.000]
22	.72044	.32328	4898.8[.000]	5105.4[.000]
23	.70897	.32849	5052.6[.000]	5272.8[.000]
24	.69671	.33345	5201.2[.000]	5435.1[.000]
25	.68432	.33817	5344.5[.000]	5592.1[.000]
26	.67190	.34267	5482.6[.000]	5744.1[.000]
27	.65914	.34695	5615.6[.000]	5890.9[.000]
28	.64617	.35102	5743.3[.000]	6032.4[.000]
29	.63190	.35488	5865.5[.000]	6168.3[.000]
30	.61782	.35854	5982.3[.000]	6298.6[.000]

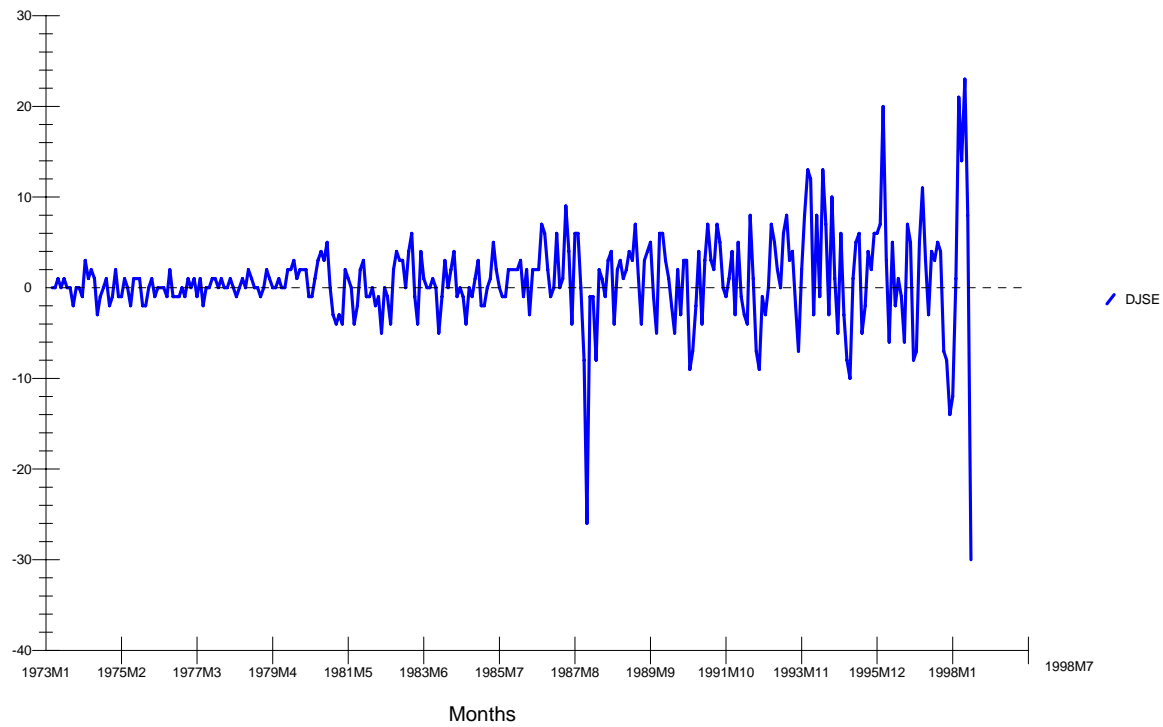
And:

Autocorrelation function of JSE, sample from 1973M2 to 1998M7



- Inference:
 - $\rho = 1$, throughout
 - JSE index nonstationary

- Now consider: First difference of JSE index

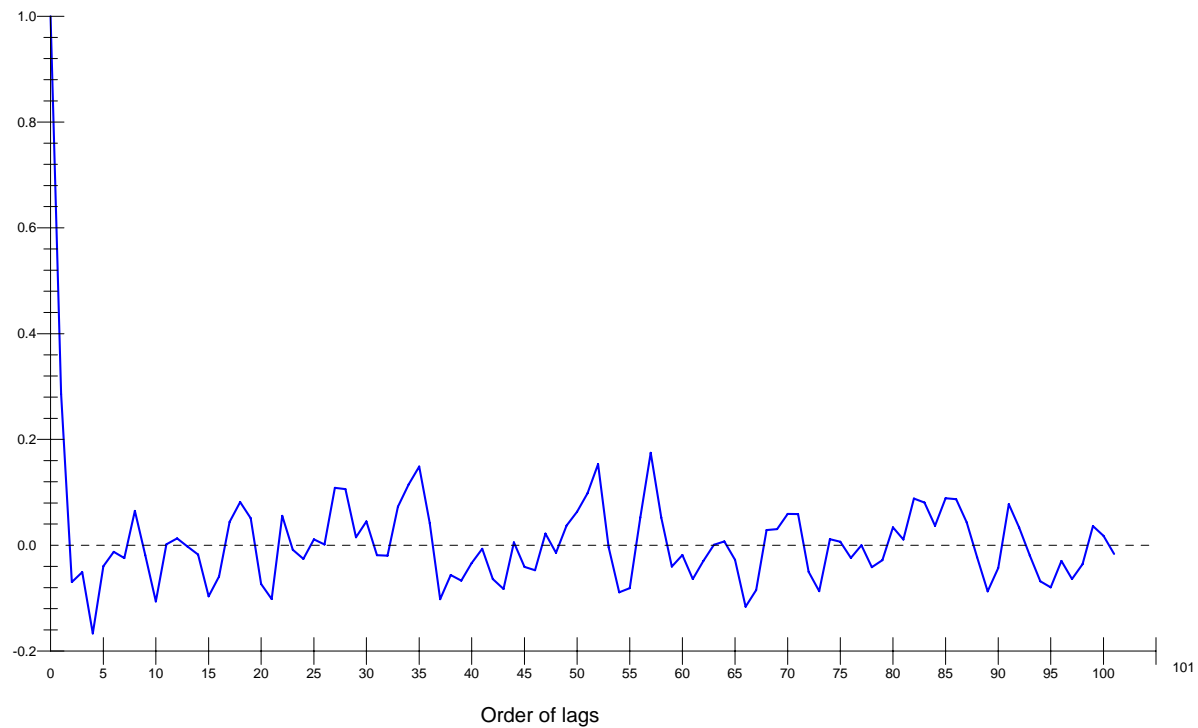


Gives:

```
Variable DJSE          Sample from 1973M3 to 1998M7
*****
Order  Autocorrelation  Standard  Box-Pierce  Ljung-Box
      Coefficient      Error    Statistic   Statistic
*****
  1      .28585          .057260   24.9218[.000]  25.1677[.000]
  2     -.069535       .061762   26.3965[.000]  26.6619[.000]
  3     -.050860       .062018   27.1855[.000]  27.4639[.000]
  4     -.16658        .062154   35.6493[.000]  36.0965[.000]
  5     -.039676       .063601   36.1295[.000]  36.5879[.000]
  6     -.012411       .063682   36.1764[.000]  36.6361[.000]
  7     -.024370       .063690   36.3576[.000]  36.8227[.000]
  8      .065229       .063721   37.6553[.000]  38.1641[.000]
  9     -.018668       .063940   37.7616[.000]  38.2744[.000]
 10     -.10617        .063957   41.1996[.000]  41.8522[.000]
 11     .0015546       .064533   41.2003[.000]  41.8530[.000]
 12     .012712       .064533   41.2496[.000]  41.9046[.000]
 13     -.0027040     .064541   41.2519[.000]  41.9070[.000]
 14     -.017792     .064541   41.3484[.000]  42.0088[.000]
 15     -.096738     .064557   44.2027[.000]  45.0304[.000]
 16     -.059125     .065031   45.2688[.000]  46.1630[.000]
 17     .043997       .065207   45.8592[.000]  46.7923[.000]
 18     .081439       .065304   47.8821[.000]  48.9562[.000]
 19     .050894       .065636   48.6721[.000]  49.8042[.000]
 20     -.073397     .065766   50.3152[.000]  51.5741[.000]
 21     -.10151       .066034   53.4578[.000]  54.9713[.000]
 22     .055598       .066543   54.4007[.000]  55.9940[.000]
 23     -.0084986     .066695   54.4227[.000]  56.0180[.000]
 24     -.025560     .066699   54.6220[.000]  56.2357[.000]
 25     .011191       .066731   54.6602[.001]  56.2776[.000]
 26     .0012694     .066737   54.6606[.001]  56.2781[.001]
 27     .10833        .066737   58.2401[.000]  60.2310[.000]
 28     .10605        .067311   61.6705[.000]  64.0329[.000]
 29     .015000       .067857   61.7391[.000]  64.1092[.000]
 30     .044885       .067868   62.3536[.000]  64.7952[.000]
```

And:

Autocorrelation function of DJSE, sample from 1973M3 to 1998M7



- Inference:
 - $\rho < 1$, throughout
 - DJSE index stationary

2.2 Spectrum

- Standardized spectral density function of a time series given by:

$$\hat{f}(\omega_j) = \lambda_0 + 2 \sum_{k=1}^m \lambda_k R_k \cos(k\omega_j)$$

$$\omega_j = \frac{j\pi}{m}$$

$$j = 1, 2, \dots, m$$

$m \equiv$ *window size*

$\lambda_k \equiv$ *set of lag weights*

$=$ *“lag window”*

- R_k denoting the autocorrelation coefficient of order k , defined by:

$$R_k = \frac{\sum_{t=k+1}^n (x_t - \bar{x})(x_{t-k} - \bar{x})}{\sum_{t=1}^n (x_t - \bar{x})^2}$$

- Alternative estimates of $\hat{f}(\omega_j)$ from alternative lag windows. Eg.'s:

$$\textit{Bartlett Window} : \lambda_k = 1 - \frac{k}{m}, \quad 0 \leq k \leq m \quad (1)$$

$$\textit{Tukey Window} : \lambda_k = \frac{1}{2} \left[1 + \cos \left(\frac{\pi k}{m} \right) \right], \quad 0 \leq k \leq m$$

$$\textit{Parzen Window} : \lambda_k = \left\{ \begin{array}{l} 1 - 6 \left(\frac{k}{m} \right)^2 + 6 \left(\frac{k}{m} \right)^3, \quad 0 \leq k \leq \frac{m}{2} \\ 2 \left(1 - \frac{k}{m} \right)^3, \quad \frac{m}{2} \leq k \leq m \end{array} \right\}$$

- Default value for $m = 2\sqrt{n}$.

- Intuition:

- Under stationarity:

- * frequency, $\hat{f}(\omega_j) \rightarrow \infty$

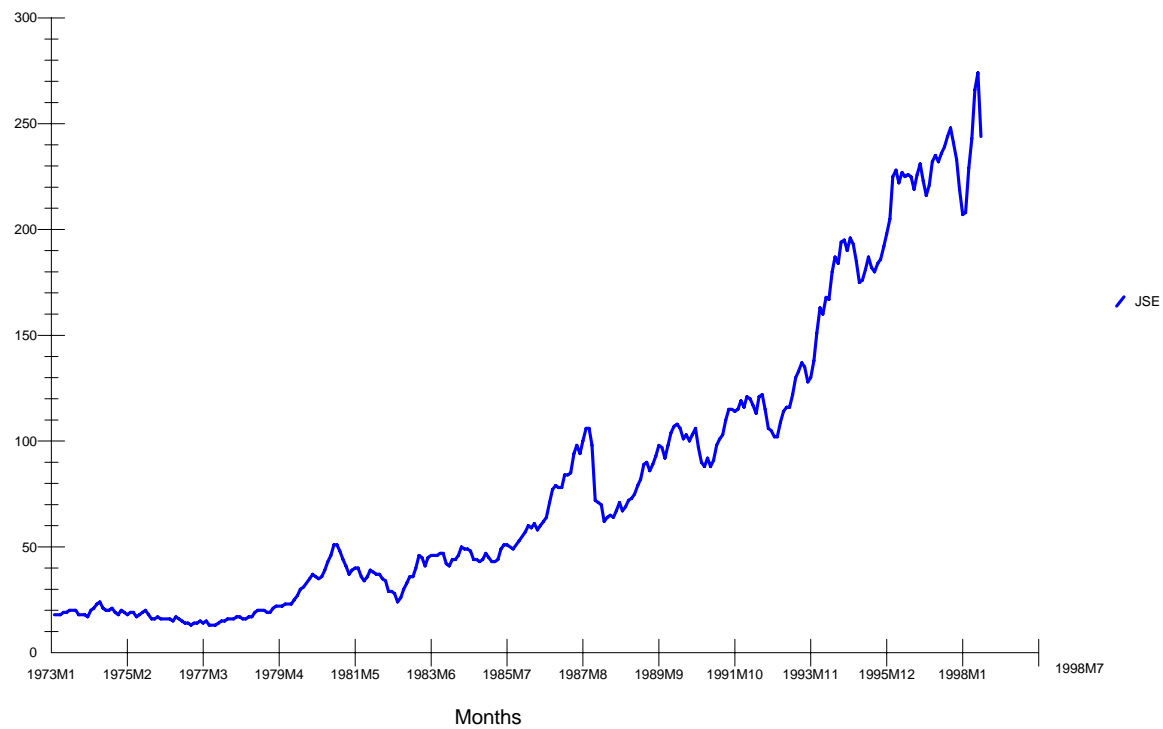
- * i.e. “frequently” see any given value of the series

- Under non-stationarity:

- * frequency, $\hat{f}(\omega_j) \rightarrow 0$

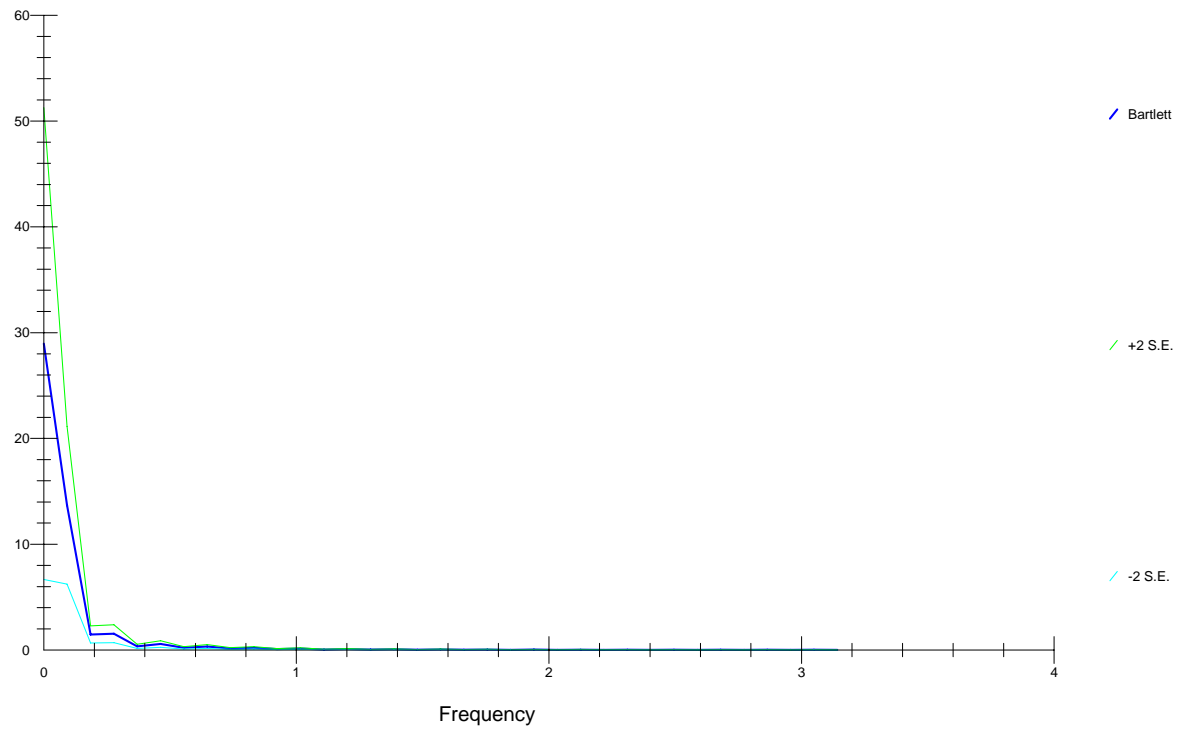
- * i.e. “infrequently” see any given value of the series

- EXAMPLE: Monthly Johannesburg Stock Exchange Index: 1973M1 1997M7



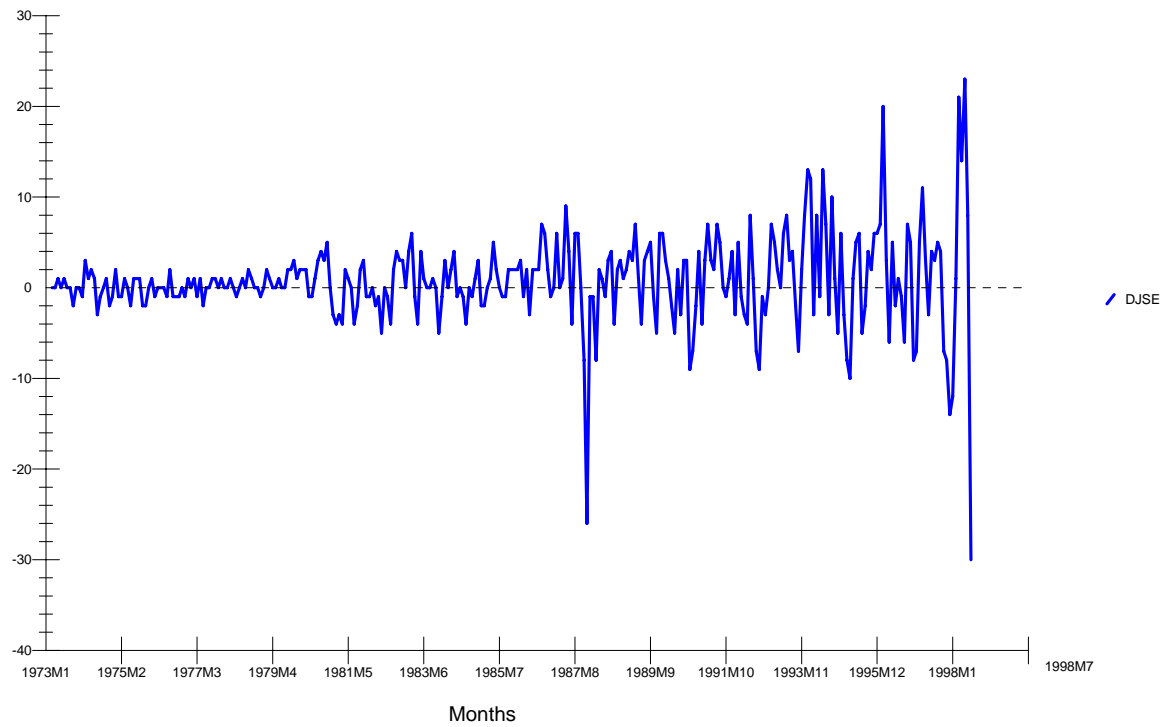
Gives:

Standardized Spectral Density Function of JSE Bartlett window

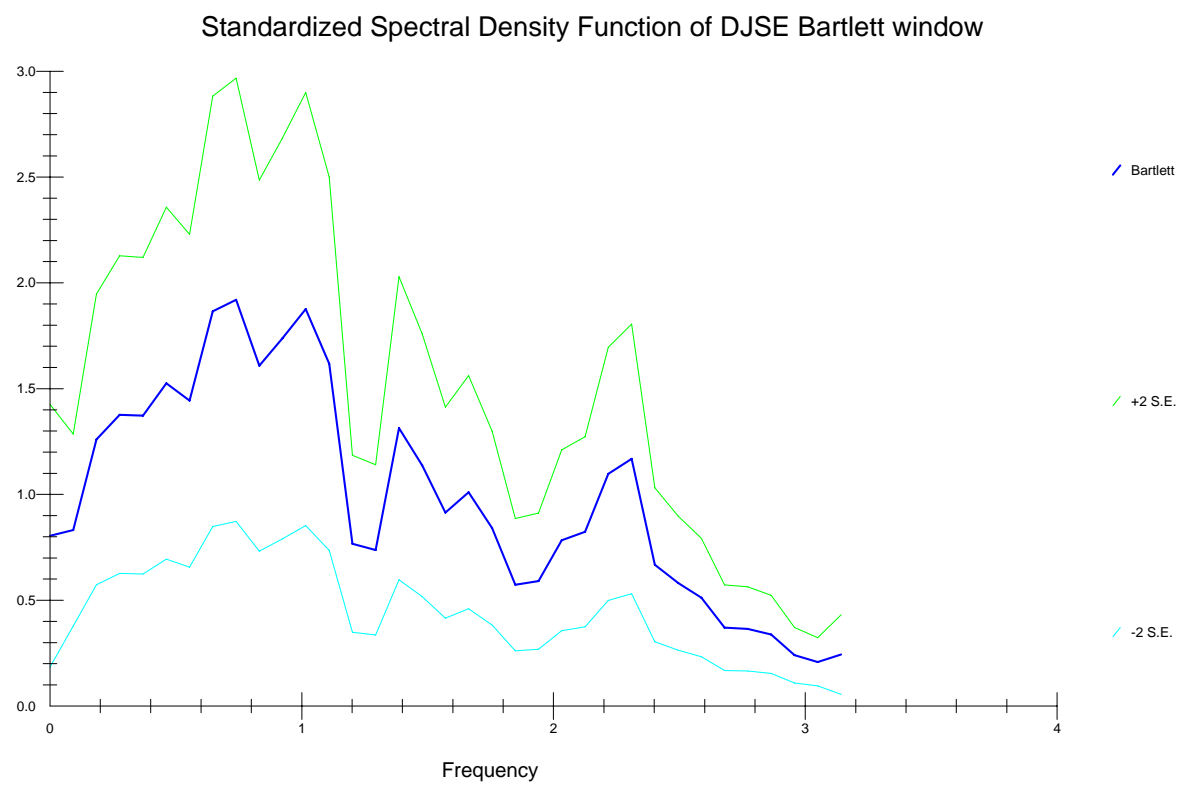


- Inference:
 - Most likely frequency $=0$
 - Hence: non-stationary

- Now consider: First difference of JSE index



Gives:



- Inference:
 - Most likely frequency >0
 - Hence: stationary

2.3 Dickey-Fuller Tests

- Most commonly used tests.
- Popularity rests on:
 - their generality
 - their simplicity
- The DF test:
 - tests directly whether $\rho = 1$
 - adjusts t-statistics: non-standard
- 3 forms of τ –test
- Plus 2 additional Φ –tests
- Requires test sequence

2.3.1 No deterministic components: no mean, no trend: τ

- The *DF*-test estimates *DGP* of the form:

$$y_t = \rho_a y_{t-1} + u_t$$

or:

$$(1 - L) y_t = \Delta y_t = (\rho_a - 1) y_{t-1} + u_t, \quad u_t \sim IID(0, \sigma^2)$$

with:

$$H_0 : \rho_a = 1$$

$$H_1 : \rho_a < 1$$

- The advantage of the second formulation, is that are testing:

$$(\rho_a - 1) = 0$$

against

$$(\rho_a - 1) < 0$$

- Critical values are from Dickey and Fuller (1981:1063).

2.3.2 Deterministic components: mean, no trend: τ_μ

- τ -form of DF test:
 - presumes DGP has zero mean,
 - and no trend:
 - \implies no deterministic components to the series.
 - Implication: initial y value is zero, since a DGP without deterministic components and a unit root has its mean determined by its initial observation.

- Where the mean of a nonstationary series is not zero, therefore, use of the initial version of the DF test would be inappropriate.
- Nankervis and Savin (1985): where the initial value of y is non-zero, τ -test leads to over-rejection of the null of nonstationarity.

- Hence τ_μ -test:

$$y_t = \mu_b + \rho_b y_{t-1} + u_t, \quad u_t \sim IID(0, \sigma^2)$$

or:

$$(1 - L) y_t = \Delta y_t = \mu_b + (\rho_b - 1) y_{t-1} + u_t, \quad u_t \sim IID(0, \sigma^2)$$

with:

$$H_0 : \rho_b = 1$$

$$H_1 : \rho_b < 1$$

- Critical values are from Dickey and Fuller (1981:1063).

2.3.3 Deterministic components: mean, trend:

τ

- Suppose DGP was a trend-stationary process:

$$y_t = \alpha + \beta t + u_t$$

- Since the DGP contains a trend component, the second version of the DF test would be capable of fitting the trend only if it set $\rho_b = 1$.
- I.e. it finds the presence of a stochastic rather than a deterministic trend.

\implies The first and second versions of the DF test effectively fail to distinguish between perhaps the two most commonly encountered time series: series with a stochastic and series with a deterministic trend.

- Hence employ version of the DF test which allows for the presence of a time trend, τ_τ -test:

$$y_t = \mu_c + \gamma_c t + \rho_c y_{t-1} + u_t, \quad u_t \sim IID(0, \sigma^2)$$

or:

$$\Delta y_t = \mu_c + \gamma_c t + (\rho_c - 1) y_{t-1} + u_t, \quad u_t \sim IID(0, \sigma^2)$$

- Note that: $\tau_\tau < \tau_\mu < \tau$,
 \implies test becomes progressively more “demanding”
- Critical values are from Dickey and Fuller (1981:1063).

- There are two points to note here, however:
 - Where τ_τ -test is inappropriately **avoided**,
 \implies **over**-accept stationarity.
 - Where τ_τ -test is inappropriately **used**,
 \implies **under**-accept stationarity.

2.3.4 Augmented Dickey-Fuller Test Statistics

- DF tests thus far: $AR(1)$ processes.
- Where DGP follows $AR(p)$ process,
 - \implies standard DF test critical values invalidated,
 - \implies error term in the tests autocorrelated,
 - \implies violate the “white noise” assumption of the DF tests.

- Correct for the possibility of autocorrelation: use *augmented DF* tests, or *ADF* tests:

$$\Delta y_t = \mu + \beta t + \vartheta_1 y_{t-1} + \sum_{i=1}^p \theta_i \Delta y_{t-i} + u_t$$

- Critical choice concerns the number of lags introduced into the *ADF* test:
 - Too few lags: over-rejection of the null of non-stationarity
 - Too many lags: reduces power of test.
- Convention: err on the side of generosity in the lag structure.

2.3.5 Dickey-Fuller Φ_3 Test: $\gamma_c = 0, \rho_c = 1$

- Testing restrictions on single coefficients by means of standard t-statistics is inappropriate in the presence of nonstationary variables.
- Testing joint restrictions by means of standard F-statistics is inappropriate in the presence of nonstationary variables.
- The Dickey-Fuller Φ_3 test tests the joint restriction that:

$$H_0 : \rho_c - 1 = \gamma_c = 0$$

- Critical values are from Dickey and Fuller (1981:1063).

2.3.6 Dickey-Fuller Φ_1 Test: $\mu_b = 0, \rho_b = 1$

- Φ_1 test tests the joint restriction that:

$$H_0 : \rho_b - 1 = \mu_b = 0$$

- Critical values are from Dickey and Fuller (1979), Table IV.

2.3.7 Perron Test Sequence

- Generalizes to *ADF* context.
- Perron suggests the following testing sequence:

1. τ_τ :

$$\Delta y_t = \mu_c + \gamma_c t + (\rho_c - 1) y_{t-1} + u_t$$

$$H_0 : \rho_c - 1 = 0, \text{ test statistic : } \tau_\tau$$

2. Φ_3 :

$$\Delta y_t = \mu_c + \gamma_c t + (\rho_c - 1) y_{t-1} + u_t$$

$$H_0 : \rho_c - 1 = \gamma_c = 0, \text{ test statistic : } \Phi_3$$

(a) t-statistic:

$$\Delta y_t = \mu_c + \gamma_c t + (\rho_c - 1) y_{t-1} + u_t$$

$$H_0 : \rho_c - 1 = 0, \text{ test statistic : } t$$

(b) For Φ_3 and Φ_1 tests, if null

$$H_0 : \rho_c = 1$$

$$H_0 : \rho_b = 1$$

not rejected, but the joint null hypotheses:

$$H_0 : \rho_c - 1 = \gamma_c = 0$$

$$H_0 : \rho_b - 1 = \mu_b = 0$$

is rejected, then:

- i. the trend/drift is significant under the null of a unit root
- ii. *asymptotic* normality of the t-statistic for ρ_i follows

3. τ_μ :

$$\Delta y_t = \mu_b + (\rho_b - 1) y_{t-1} + u_t$$

$$H_0 : \rho_b - 1 = 0, \text{ test statistic : } \tau_\mu$$

4. Φ_1 :

$$\Delta y_t = \mu_b + (\rho_b - 1) y_{t-1} + u_t$$

$$H_0 : \rho_b - 1 = \mu_b = 0, \text{ test statistic : } \Phi_1$$

(a) t-statistic:

$$\Delta y_t = \mu_b + (\rho_b - 1) y_{t-1} + u_t$$

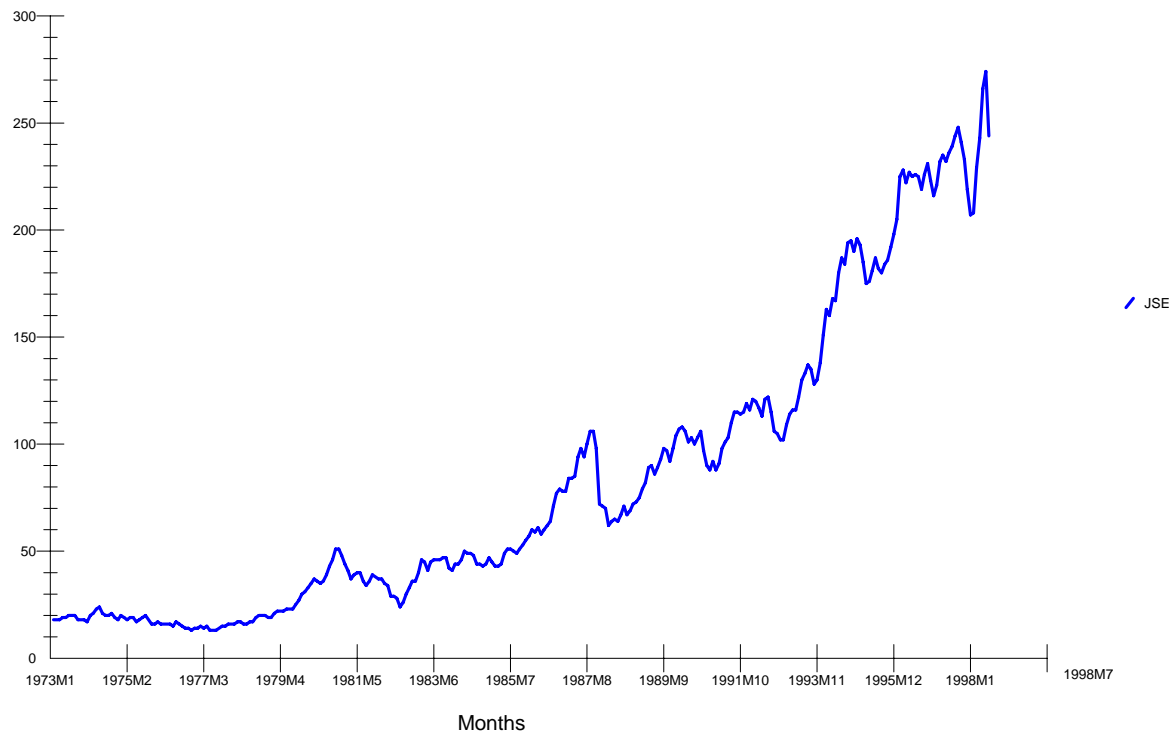
$$H_0 : \rho_b - 1 = 0, \text{ test statistic : } t$$

5. τ :

$$\Delta y_t = (\rho_a - 1) y_{t-1} + u_t$$

$$H_0 : \rho_a - 1 = 0, \text{ test statistic : } \tau$$

- EXAMPLE: Monthly Johannesburg Stock Exchange Index: 1973M1 1997M7



- Gives τ_T :

```

Unit root tests for variable JSE
The Dickey-Fuller regressions include an intercept and a linear trend
*****
293 observations used in the estimation of all ADF regressions.
Sample period from 1974M3 to 1998M7
*****
Test Statistic      LL          AIC          SBC          HQC
DF                  -1.5333     -888.6219    -891.6219    -897.1422    -893.8329
ADF(1)              -2.4641     -873.4903    -877.4903    -884.8506    -880.4382
ADF(2)              -1.9207     -867.5362    -872.5362    -881.7367    -876.2211
ADF(3)              -1.9334     -867.5049    -873.5049    -884.5454    -877.9268
ADF(4)              -1.5580     -860.3410    -867.3410    -880.2216    -872.4998
ADF(5)              -1.6303     -859.6836    -867.6836    -882.4043    -873.5794
ADF(6)              -1.5315     -858.9270    -867.9270    -884.4878    -874.5598
ADF(7)              -1.5624     -858.8211    -868.8211    -887.2220    -876.1909
ADF(8)              -1.7328     -857.5435    -868.5435    -888.7845    -876.6503
ADF(9)              -1.5399     -856.4118    -868.4118    -890.4928    -877.2555
ADF(10)             -1.3354     -855.4353    -868.4353    -892.3565    -878.0161
ADF(11)             -1.3724     -855.3549    -869.3549    -895.1161    -879.6727
ADF(12)             -1.3043     -855.2338    -870.2338    -897.8351    -881.2885
*****
95% critical value for the augmented Dickey-Fuller statistic = -3.4266
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

- Inference: $H_0 : \rho_c - 1 = 0$ is not rejected,

• And Φ_3 :

```

Ordinary Least Squares Estimation
*****
Dependent variable is DJSE
301 observations used for estimation from 1973M7 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
C              -.64842          .62177              -1.0429[.298]
T              .016307         .0077096           2.1152[.035]
JSE(-1)       -.014581         .010035            -1.4530[.147]
DJSE(-1)      .41530          .061405            6.7633[.000]
DJSE(-2)     -.26168          .066133            -3.9569[.000]
DJSE(-3)      .099438         .067373            1.4759[.141]
DJSE(-4)     -.24655          .063867            -3.8604[.000]
*****
R-Squared      .19072          R-Bar-Squared      .17421
S.E. of Regression  4.5620      F-stat.      F( 6, 294)  11.5479[.000]
Mean of Dependent Variable  .74751      S.D. of Dependent Variable  5.0202
Residual Sum of Squares  6118.8      Equation Log-likelihood  -880.4078
Akaike Info. Criterion  -887.4078      Schwarz Bayesian Criterion  -900.3827
DW-statistic    1.8779
*****

Diagnostic Tests
*****
* Test Statistics *      LM Version      *      F Version      *
*****
* A:Serial Correlation*CHSQ( 12)= 13.1649[.357]*F( 12, 282)= 1.0748[.381]*
*
* B:Functional Form *CHSQ( 1)= .057116[.811]*F( 1, 293)= .055609[.814]*
*
* C:Normality *CHSQ( 2)= 522.8766[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 12.7291[.000]*F( 1, 299)= 13.2029[.000]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

```

Variable Deletion Test (OLS case)
*****
Dependent variable is DJSE
List of the variables deleted from the regression:
T
JSE(-1)
301 observations used for estimation from 1973M7 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
C              .70617           .27475              2.5703[.011]
DJSE(-1)      .42063           .061045            6.8905[.000]
DJSE(-2)     -.26458           .065908            -4.0143[.000]
DJSE(-3)      .10046           .067519            1.4879[.138]
DJSE(-4)     -.24553           .063688            -3.8552[.000]
*****
Joint test of zero restrictions on the coefficients of deleted variables:
Lagrange Multiplier Statistic   CHSQ( 2)= 5.9316[.052]
Likelihood Ratio Statistic       CHSQ( 2)= 5.9909[.050]
F Statistic                       F( 2, 294)= 2.9551[.054]
*****

```

- Critical value: ≈ 6.34

- Inference:

- $\rho_c - 1 = \gamma_c = 0$, not rejected

- Hence:

- * the trend is *not* significant under the null of a unit root

- * *asymptotic* normality of the t-statistic for ρ_i does not follow

- And τ_μ :

```

Unit root tests for variable JSE
The Dickey-Fuller regressions include an intercept but not a trend
*****
293 observations used in the estimation of all ADF regressions.
Sample period from 1974M3 to 1998M7
*****
Test Statistic      LL          AIC          SBC          HQC
DF                   1.1515      -890.9399    -892.9399    -896.6200    -894.4138
ADF(1)               .10451      -877.1988    -880.1988    -885.7191    -882.4098
ADF(2)               .67985      -870.3829    -874.3829    -881.7432    -877.3308
ADF(3)               .66204      -870.3795    -875.3795    -884.5799    -879.0644
ADF(4)               1.1445      -862.7645    -868.7645    -879.8050    -873.1864
ADF(5)               1.0182      -862.1776    -869.1776    -882.0582    -874.3365
ADF(6)               1.1594      -861.3330    -869.3330    -884.0536    -875.2288
ADF(7)               1.0949      -861.2597    -870.2597    -886.8205    -876.8926
ADF(8)               .84352      -860.1772    -870.1772    -888.5780    -877.5470
ADF(9)               1.0998      -858.8446    -869.8446    -890.0856    -877.9514
ADF(10)              1.3762      -857.6770    -869.6770    -891.7580    -878.5207
ADF(11)              1.2923      -857.6354    -870.6354    -894.5566    -880.2162
ADF(12)              1.3729      -857.4605    -871.4605    -897.2217    -881.7783
*****
95% critical value for the augmented Dickey-Fuller statistic = -2.8715
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

- Inference: $H_0 : \rho_b - 1 = 0$ is not rejected,

• And Φ_1 :

```

Ordinary Least Squares Estimation
*****
Dependent variable is DJSE
301 observations used for estimation from 1973M7 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
C              .33500           .41526              .80671[.420]
JSE(-1)       .0048459        .0040676           1.1914[.234]
DJSE(-1)      .40956          .061705            6.6375[.000]
DJSE(-2)     -.27348         .066285           -4.1259[.000]
DJSE(-3)      .093647         .067713            1.3830[.168]
DJSE(-4)     -.25484         .064121           -3.9744[.000]
*****
R-Squared      .17841          R-Bar-Squared      .16448
S.E. of Regression  4.5888      F-stat.      F( 5, 295)  12.8118[.000]
Mean of Dependent Variable  .74751      S.D. of Dependent Variable  5.0202
Residual Sum of Squares  6211.9      Equation Log-likelihood  -882.6809
Akaike Info. Criterion  -888.6809   Schwarz Bayesian Criterion  -899.8022
DW-statistic    1.8744
*****

```

```

Diagnostic Tests
*****
* Test Statistics *      LM Version      *      F Version      *
*****
* A:Serial Correlation*CHSQ( 12)= 14.5657[.266]*F( 12, 283)= 1.1993[.283]*
*
* B:Functional Form *CHSQ( 1)= .12887[.720]*F( 1, 294)= .12592[.723]*
*
* C:Normality *CHSQ( 2)= 514.4048[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 11.0766[.001]*F( 1, 299)= 11.4234[.001]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

```

Variable Deletion Test (OLS case)
*****
Dependent variable is DJSE
List of the variables deleted from the regression:
C          JSE(-1)
301 observations used for estimation from 1973M7 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
DJSE(-1)       .44091            .061102             7.2160[.000]
DJSE(-2)      -.25177           .066337            -3.7953[.000]
DJSE(-3)       .11351           .067960            1.6703[.096]
DJSE(-4)      -.22656           .063853            -3.5481[.000]
*****
Joint test of zero restrictions on the coefficients of deleted variables:
Lagrange Multiplier Statistic    CHSQ( 2)= 7.9810[.018]
Likelihood Ratio Statistic        CHSQ( 2)= 8.0887[.018]
F Statistic                        F( 2, 295)= 4.0175[.019]
*****

```

- Critical value: ≈ 4.63
- Inference:
 - $\rho_b - 1 = \mu_b = 0$, not rejected
 - Hence:
 - * the drift is *not* significant under the null of a unit root
 - * *asymptotic* normality of the t-statistic for ρ_i does not follow

- And τ :

```

                          Ordinary Least Squares Estimation
*****
Dependent variable is DJSE
301 observations used for estimation from 1973M7 to 1998M7
*****
Regressor           Coefficient           Standard Error           T-Ratio[Prob]
JSE(-1)             .0073078             .0026877                 2.7190[.007]
DJSE(-1)            .40815               .061644                  6.6211[.000]
DJSE(-2)            -.27535              .066205                  -4.1591[.000]
DJSE(-3)            .092890              .067667                  1.3728[.171]
DJSE(-4)            -.25564              .064076                  -3.9897[.000]
*****
R-Squared           .17660               R-Bar-Squared           .16547
S.E. of Regression  4.5861               F-stat. F( 4, 296)     15.8707[.000]
Mean of Dependent Variable .74751           S.D. of Dependent Variable 5.0202
Residual Sum of Squares 6225.6           Equation Log-likelihood -883.0126
Akaike Info. Criterion -888.0126        Schwarz Bayesian Criterion -897.2803
DW-statistic        1.8724
*****

```

```

                          Diagnostic Tests
*****
*   Test Statistics   *   LM Version   *   F Version   *
*****
*   A:Serial Correlation*CHSQ( 12)= 14.6988[.258]*F( 12, 284)= 1.2151[.272]*
*   *               *   *               *   *               *
*   B:Functional Form *CHSQ( 1)= .23549[.627]*F( 1, 295)= .23097[.631]*
*   *               *   *               *   *               *
*   C:Normality      *CHSQ( 2)= 516.1714[.000]*   Not applicable   *
*   *               *   *               *   *               *
*   D:Heteroscedasticity*CHSQ( 1)= 10.4869[.001]*F( 1, 299)= 10.7933[.001]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

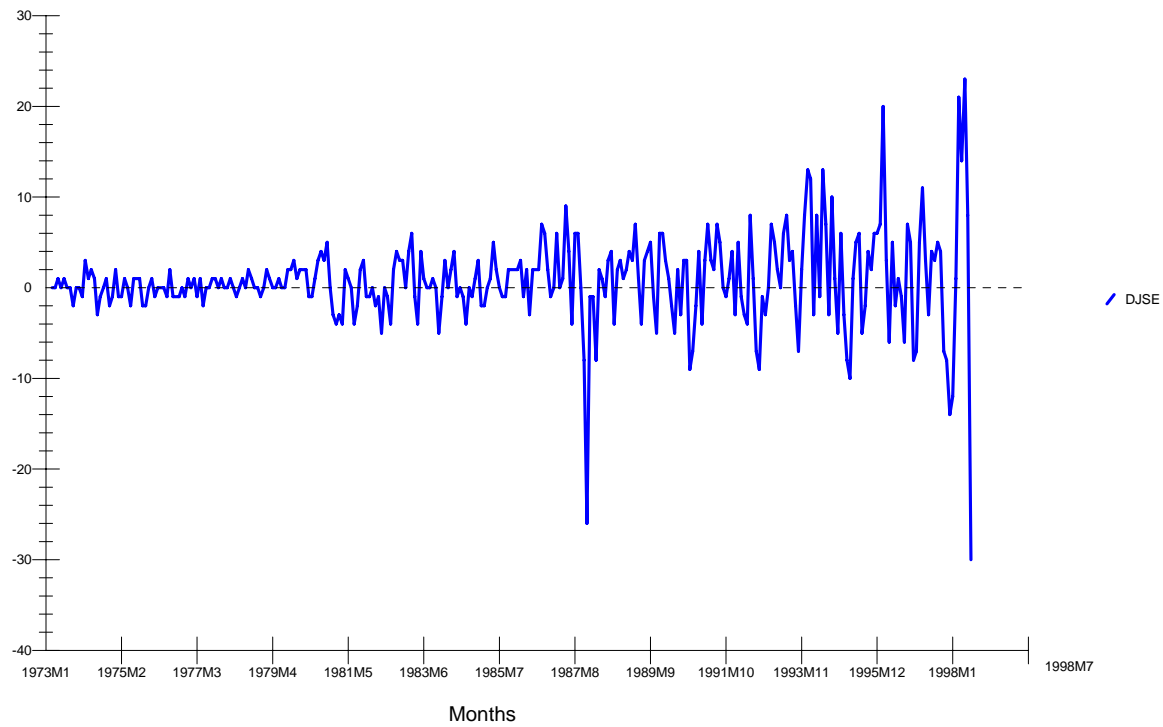
```

- Inference:

- $\rho_a - 1 = 0$, not rejected

- JSE nonstationary.

- Now consider: First difference of JSE index



- Gives τ_{τ} :

```

Unit root tests for variable DJSE
The Dickey-Fuller regressions include an intercept and a linear trend
*****
292 observations used in the estimation of all ADF regressions.
Sample period from 1974M4 to 1998M7
*****
Test Statistic      LL          AIC          SBC          HQC
DF                  -11.2958    -874.0434    -877.0434    -882.5585    -879.2525
ADF(1)             -11.6444    -866.9308    -870.9308    -878.2843    -873.8763
ADF(2)             -9.5632     -866.9305    -871.9305    -881.1224    -875.6124
ADF(3)             -10.4857    -859.1350    -865.1350    -876.1653    -869.5533
ADF(4)             -8.3173     -858.6026    -865.6026    -878.4712    -870.7572
ADF(5)             -8.0488     -857.6947    -865.6947    -880.4017    -871.5858
ADF(6)             -7.1139     -857.6423    -866.6423    -883.1877    -873.2697
ADF(7)             -5.9863     -856.6592    -866.6592    -885.0430    -874.0230
ADF(8)             -6.2128     -855.2108    -866.2108    -886.4330    -874.3110
ADF(9)             -6.3615     -853.9412    -865.9412    -888.0018    -874.7778
ADF(10)            -5.8121     -853.9165    -866.9165    -890.8154    -876.4895
ADF(11)            -5.6796     -853.7043    -867.7043    -893.4416    -878.0136
ADF(12)            -5.6130     -853.3948    -868.3948    -895.9704    -879.4405
*****
95% critical value for the augmented Dickey-Fuller statistic = -3.4266
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

- Inference: $H_0 : \rho_c - 1 = 0$ is rejected,

• And Φ_3 :

```

Ordinary Least Squares Estimation
*****
Dependent variable is DDJSE
300 observations used for estimation from 1973M8 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
C              -.22516          .54974              -.40958[.682]
T              .0058535        .0031462            1.8605[.064]
DJSE(-1)      -.96827         .11472              -8.4403[.000]
DDJSE(-1)     .38390         .10249              3.7458[.000]
DDJSE(-2)     .10414         .095354             1.0922[.276]
DDJSE(-3)     .21118         .077425             2.7275[.007]
DDJSE(-4)     -.068543        .066674             -1.0280[.305]
*****
R-Squared      .37679          R-Bar-Squared      .36403
S.E. of Regression  4.5772        F-stat.      F( 6, 293)  29.5246[.000]
Mean of Dependent Variable  -.10333      S.D. of Dependent Variable  5.7396
Residual Sum of Squares  6138.5      Equation Log-likelihood  -878.4639
Akaike Info. Criterion  -885.4639   Schwarz Bayesian Criterion  -898.4271
DW-statistic   1.8954
*****

```

```

Diagnostic Tests
*****
* Test Statistics *      LM Version      *      F Version      *
*****
* A:Serial Correlation*CHSQ( 12)= 12.3098[.421]*F( 12, 281)= 1.0020[.447]*
*
* B:Functional Form *CHSQ( 1)= 3.9031[.048]*F( 1, 292)= 3.8491[.051]*
*
* C:Normality *CHSQ( 2)= 486.1393[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 31.0233[.000]*F( 1, 298)= 34.3708[.000]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

```

Variable Deletion Test (OLS case)
*****
Dependent variable is DDJSE
List of the variables deleted from the regression:
C          T
300 observations used for estimation from 1973M8 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
DJSE(-1)       -.83884           .10768              -7.7904[.000]
DDJSE(-1)      .29096           .098925             2.9412[.004]
DDJSE(-2)      .028381          .093092             .30487[.761]
DDJSE(-3)      .16183           .076584             2.1131[.035]
DDJSE(-4)      -.10174          .066544             -1.5289[.127]
*****
Joint test of zero restrictions on the coefficients of deleted variables:
Lagrange Multiplier Statistic   CHSQ( 2)= 9.0484[.011]
Likelihood Ratio Statistic       CHSQ( 2)= 9.1876[.010]
F Statistic                       F( 2, 293)= 4.5560[.011]
*****

```

- Critical value: ≈ 6.34

- Inference:

- $\rho_c - 1 = \gamma_c = 0$, not rejected

- Hence:

- * the trend is *not* significant under the null of a unit root

- * *asymptotic* normality of the t-statistic for ρ_i does not follow

- And τ_μ :

```

Unit root tests for variable DJSE
The Dickey-Fuller regressions include an intercept but not a trend
*****
292 observations used in the estimation of all ADF regressions.
Sample period from 1974M4 to 1998M7
*****
Test Statistic      LL          AIC          SBC          HQC
DF                  -11.2497    -874.7051    -876.7051    -880.3818    -878.1778
ADF(1)             -11.5324    -868.1196    -871.1196    -876.6347    -873.3287
ADF(2)              -9.4301     -868.1048    -872.1048    -879.4583    -875.0503
ADF(3)             -10.2764    -860.9471    -865.9471    -875.1390    -869.6290
ADF(4)              -8.0926     -860.2349    -866.2349    -877.2651    -870.6531
ADF(5)              -7.7844     -859.5514    -866.5514    -879.4201    -871.7061
ADF(6)              -6.8388     -859.4074    -867.4074    -882.1145    -873.2985
ADF(7)              -5.7362     -858.0756    -867.0756    -883.6210    -873.7030
ADF(8)              -5.9040     -856.9973    -866.9973    -885.3811    -874.3611
ADF(9)              -5.9762     -856.1960    -867.1960    -887.4181    -875.2962
ADF(10)             -5.4227     -856.0449    -868.0449    -890.1055    -876.8815
ADF(11)             -5.2536     -855.9844    -868.9844    -892.8833    -878.5573
ADF(12)             -5.1442     -855.8714    -869.8714    -895.6087    -880.1807
*****
95% critical value for the augmented Dickey-Fuller statistic = -2.8716
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

- Inference: $H_0 : \rho_b - 1 = 0$ is rejected,

• And Φ_1 :

```

Ordinary Least Squares Estimation
*****
Dependent variable is DDJSE
300 observations used for estimation from 1973M8 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
C              .65834           .27811              2.3672[.019]
DJSE(-1)      -.91762          .11191              -8.1998[.000]
DDJSE(-1)     .34842           .10112              3.4457[.001]
DDJSE(-2)     .076283          .094564             .80668[.421]
DDJSE(-3)     .19329           .077147             2.5055[.013]
DDJSE(-4)     -.080081         .066662             -1.2013[.231]
*****
R-Squared      .36943           R-Bar-Squared      .35871
S.E. of Regression  4.5963         F-stat.      F( 5, 294)  34.4489[.000]
Mean of Dependent Variable  -.10333       S.D. of Dependent Variable  5.7396
Residual Sum of Squares  6211.0        Equation Log-likelihood  -880.2256
Akaike Info. Criterion  -886.2256     Schwarz Bayesian Criterion  -897.3369
DW-statistic   1.8987
*****

```

```

Diagnostic Tests
*****
* Test Statistics *      LM Version      *      F Version      *
*****
* A:Serial Correlation*CHSQ( 12)= 10.1276[.605]*F( 12, 282)= .82104[.629]*
*
* B:Functional Form *CHSQ( 1)= 1.1845[.276]*F( 1, 293)= 1.1615[.282]*
*
* C:Normality *CHSQ( 2)= 471.9442[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 32.5907[.000]*F( 1, 298)= 36.3190[.000]*
*****

```

```

A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

Variable Deletion Test (OLS case)

Dependent variable is DDJSE

List of the variables deleted from the regression:

C DJSE(-1)

300 observations used for estimation from 1973M8 to 1998M7

Regressor	Coefficient	Standard Error	T-Ratio[Prob]
DDJSE(-1)	-.34569	.061110	-5.6568[.000]
DDJSE(-2)	-.52586	.065814	-7.9900[.000]
DDJSE(-3)	-.21112	.065528	-3.2218[.001]
DDJSE(-4)	-.36990	.062428	-5.9252[.000]

Joint test of zero restrictions on the coefficients of deleted variables:

Lagrange Multiplier Statistic CHSQ(2)= 55.8420[.000]

Likelihood Ratio Statistic CHSQ(2)= 61.7901[.000]

F Statistic F(2, 294)= 33.6208[.000]

- Critical value: ≈ 4.63
- Inference:
 - $\rho_b - 1 = \mu_b = 0$, is rejected
 - Hence:
 - * the drift *is* significant under the null of a unit root
 - \implies invalidates τ test
 - * *asymptotic* normality of the t-statistic for ρ_i does follow

● So:

```
Ordinary Least Squares Estimation
*****
Dependent variable is DDJSE
300 observations used for estimation from 1973M8 to 1998M7
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                   .65834           .27811              2.3672[.019]
DJSE(-1)           -.91762          .11191              -8.1998[.000]
DDJSE(-1)          .34842           .10112              3.4457[.001]
DDJSE(-2)          .076283          .094564             .80668[.421]
DDJSE(-3)          .19329           .077147             2.5055[.013]
DDJSE(-4)          -.080081         .066662             -1.2013[.231]
*****
R-Squared           .36943           R-Bar-Squared       .35871
S.E. of Regression  4.5963           F-stat.             F( 5, 294)         34.4489[.000]
Mean of Dependent Variable -.10333         S.D. of Dependent Variable  5.7396
Residual Sum of Squares  6211.0          Equation Log-likelihood -880.2256
Akaike Info. Criterion  -886.2256       Schwarz Bayesian Criterion -897.3369
DW-statistic        1.8987
*****
```

```
Diagnostic Tests
*****
* Test Statistics *          LM Version          *          F Version          *
*****
* A:Serial Correlation*CHSQ( 12)= 10.1276[.605]*F( 12, 282)= .82104[.629]*
*
* B:Functional Form *CHSQ( 1)= 1.1845[.276]*F( 1, 293)= 1.1615[.282]*
*
* C:Normality *CHSQ( 2)= 471.9442[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 32.5907[.000]*F( 1, 298)= 36.3190[.000]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values
```

- Inference:

- $-0.91762 < 0$

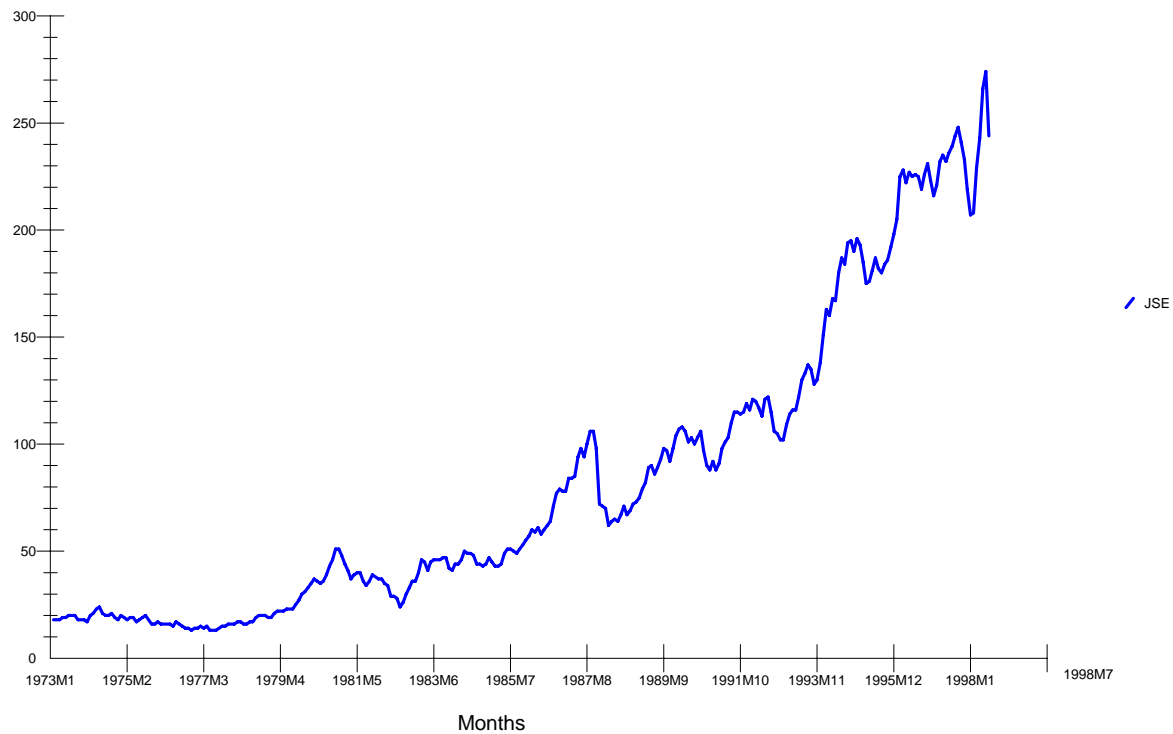
- $\implies \rho < 1$

- Stationary, subject to drift.

2.4 Phillips-Perron Test

- Alternative to ADF test
- Proposed by Phillips (1987), Perron (1988) and Phillips and Perron (1988).
- Standard DF tests.
- Undertaking a nonparametric correction to the t-test to “compensate” for the biasing effects of the autocorrelation.
- Critical values: ADF .
- Monte Carlo work (see Schwert 1989):
 - poor size properties,
 - reject null of nonstationarity too easily when DGP has large negative MA components.

- EXAMPLE: Monthly Johannesburg Stock Exchange Index: 1973M1 1997M7



Gives:

```
Ordinary Least Squares Estimation
*****
Dependent variable is DJSE
305 observations used for estimation from 1973M3 to 1998M7
*****
Regressor      Coefficient      Standard Error      T-Ratio[Prob]
C              .32241          .44456              .72522[.469]
JSE(-1)       .0051296       .0041777           1.2279[.220]
*****
R-Squared      .0049511      R-Bar-Squared      .0016671
S.E. of Regression  4.9835      F-stat.      F( 1, 303)      1.5077[.220]
Mean of Dependent Variable  .74098      S.D. of Dependent Variable  4.9877
Residual Sum of Squares  7525.1      Equation Log-likelihood  -921.6435
Akaike Info. Criterion  -923.6435      Schwarz Bayesian Criterion  -927.3638
DW-statistic  1.3159
*****
```

```
Diagnostic Tests
*****
*      Test Statistics      *      LM Version      *      F Version      *
*****
*      *      *      *      *
* A:Serial Correlation*CHSQ( 12)= 62.0648[.000]*F( 12, 291)= 6.1954[.000]*
*      *      *      *      *
* B:Functional Form      *CHSQ( 1)= 3.7203[.054]*F( 1, 302)= 3.7292[.054]*
*      *      *      *      *
* C:Normality          *CHSQ( 2)= 1224.2[.000]*      Not applicable      *
*      *      *      *      *
* D:Heteroscedasticity*CHSQ( 1)= 49.0522[.000]*F( 1, 303)= 58.0698[.000]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values
```

And then:

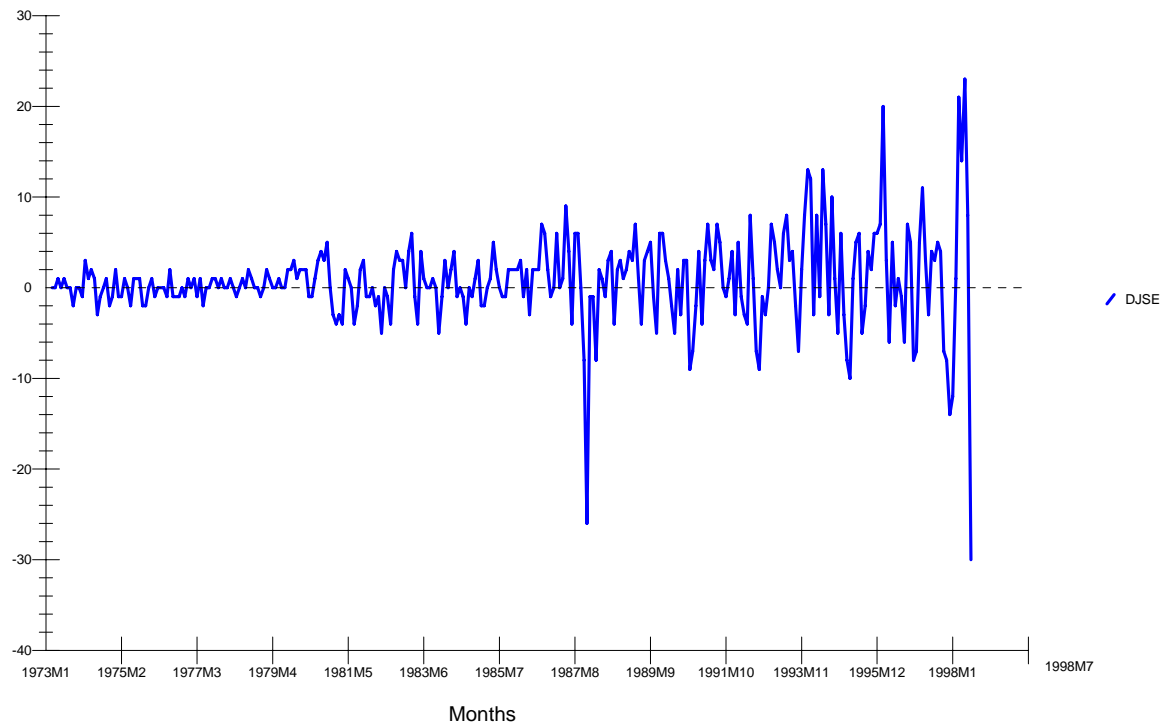
```
                Ordinary Least Squares Estimation
Based on Newey-West adjusted S.E.'s Bartlett weights, truncation lag= 12
*****
Dependent variable is DJSE
305 observations used for estimation from 1973M3 to 1998M7
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                  .32241             .30000              1.0747[.283]
JSE(-1)           .0051296          .0043863            1.1695[.243]
*****
```

- Inference:

- $\rho = 1$

- Non-stationary

- Now consider: First difference of JSE index



Gives:

```
Ordinary Least Squares Estimation
*****
Dependent variable is DDJSE
304 observations used for estimation from 1973M4 to 1998M7
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                  .46821          .27768              1.6861[.093]
DJSE(-1)          -.67319         .058583             -11.4912[.000]
*****
R-Squared          .30422          R-Bar-Squared       .30192
S.E. of Regression 4.7645          F-stat.             F( 1, 302) 132.0471[.000]
Mean of Dependent Variable -.098684        S.D. of Dependent Variable 5.7025
Residual Sum of Squares 6855.5          Equation Log-likelihood -904.9561
Akaike Info. Criterion -906.9561        Schwarz Bayesian Criterion -910.6731
DW-statistic       1.7160
*****
```

```
Diagnostic Tests
*****
* Test Statistics * LM Version * F Version *
*****
* A:Serial Correlation*CHSQ( 12)= 36.2728[.000]*F( 12, 290)= 3.2742[.000]*
* * * * *
* B:Functional Form *CHSQ( 1)= 1.3857[.239]*F( 1, 301)= 1.3783[.241]*
* * * * *
* C:Normality *CHSQ( 2)= 1464.8[.000]* Not applicable *
* * * * *
* D:Heteroscedasticity*CHSQ( 1)= 4.4829[.034]*F( 1, 302)= 4.5201[.034]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values
```

And then:

```
                Ordinary Least Squares Estimation
Based on Newey-West adjusted S.E.'s Bartlett weights, truncation lag= 12
*****
Dependent variable is DDJSE
304 observations used for estimation from 1973M4 to 1998M7
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                   .46821             .21258              2.2025[.028]
DJSE(-1)           -.67319            .065441            -10.2870[.000]
*****
```

- Inference:

- $\rho < 1$

- Stationary

2.5 Perron Test: Structural Breaks

- Perron (1989): trend-stationary series with a structural break,
 - that is a series which has undergone a permanent change either in slope or intercept at some point during the time period being tested for stationarity,
 - would be considered as a persistent innovation in a series with stochastic trend by the usual ADF -statistic
 - See also Holden and Perman (1994).
 - Thus the test would show nonstationarity, where in fact stationarity applies.

- Alternative responses:
 - Timing of break is known
 - Timing of break is not known

1. Where the time point at which the break occurs in the series is known:
 - (a) Include dummy variables into the deterministic component of the relevant DF or ADF statistic, to correct for the intercept or slope change of the series.
 - i. Critical values for the unit root tests are found in Perron (1989, 1990).

(b) Conduct a further set of tests proposed by Perron (1994):

i. Innovational outlier model: where structural breaks do not have an instantaneous effect:

$$y_t = \mu + \beta t + \theta DU_t + \gamma DT_t + \delta DTB_t + \alpha y_{t-1} + \sum \alpha_i \Delta y_{t-i} + e_t$$

$$H_0 : \alpha = 1$$

$$DU_t = 1, DT_t = t - T_b \text{ if } t > T_b, 0 \text{ otherwise}$$

$$DTB_t = 1 \text{ if } t = T_b + 1, 0 \text{ otherwise}$$

where T_b denotes the time point of structural break.

ii. Additive outlier model where the effect of the break is instantaneous:

$$y_t^* = \alpha y_{t-1}^* + \sum d_j DTB_{t-j} + \sum \alpha_i \Delta y_{t-i}^* + e_t$$

$$H_0 : \alpha = 1$$

$$y_t^* = y_t - \mu - \beta t - \gamma DT_t$$

$$DTB_t = 1 \text{ if } t = T_b + 1, 0 \text{ otherwise}$$

iii. Critical values are to be found in Perron (1994: Tables 4.2 to 4.5).

2. The time point for a structural break is unknown.

⇒ Resort to recursive, rolling or sequential tests for unit roots:

(a) The usual DF or ADF statistics are computed over various sub-samples of the data (not entirely reliable).

- (b) Recursive tests: the usual DF or ADF statistics are computed over subsamples $t = 1, \dots, k$, for $k = k_1, \dots, T$, where k_1 is a start-up value, T full sample size. The minimum value of $\tau_\tau(k/T)$ obtained from all subsamples is chosen and compared to the critical values listed in Banerjee, Lumsdaine and Stock (1992) Table 1, to test for the null of a unit root.
- (c) Rolling tests: proceed as for recursive tests, except that the sub-samples are a constant fraction of the total sample, and the fraction over which estimation takes place “rolls” through the sample.

(d) Sequential tests: use an adaptation of the *ADF* test:

$$\Delta y_t = \mu + \beta_1 t + \beta_2 D + \vartheta_1 y_{t-i} + \sum_{i=1}^p \theta_i \Delta y_{t-i} + u_t \quad (2)$$

which allows for the shift in trend model where:

$$\begin{aligned} D &= t \text{ for } t > k \\ D &= 0 \text{ for } t \leq k \end{aligned} \quad (3)$$

and the shift in mean model where:

$$\begin{aligned} D &= 1 \text{ for } t > k \\ D &= 0 \text{ for } t \leq k \end{aligned} \quad (4)$$

But k is unknown;

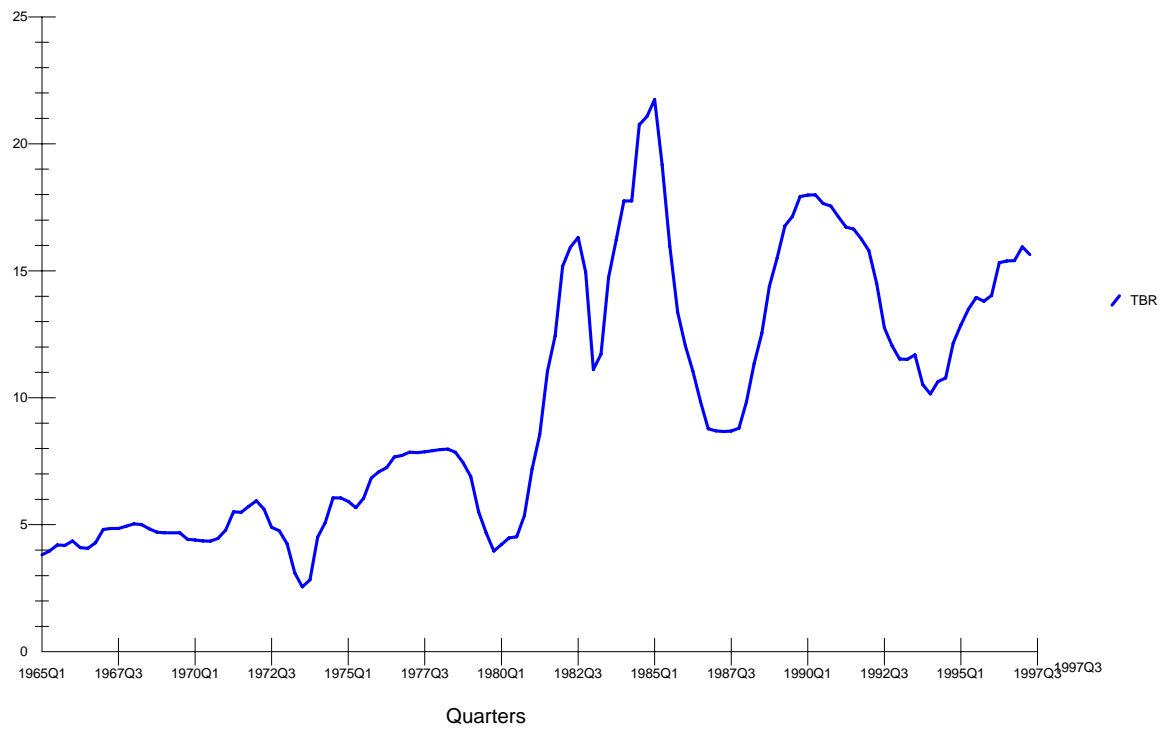
\implies Banerjee, Lumsdaine and Stock (1992) suggest searching through $0.15T \leq k \leq T - 0.15T$.

\implies The minimum value of $\tau_\tau(k/T)$ from the search is compared to the critical values listed in Banerjee, Lumsdaine and Stock (1992) Table 2, to test for the null of a unit root.

\implies We may also use the F-test of $H_0 : \beta_2 = \vartheta_1 = 0$, and compare the largest sequentially calculated F -value to the critical values listed in Banerjee, Lumsdaine and Stock (1992) Table 2.

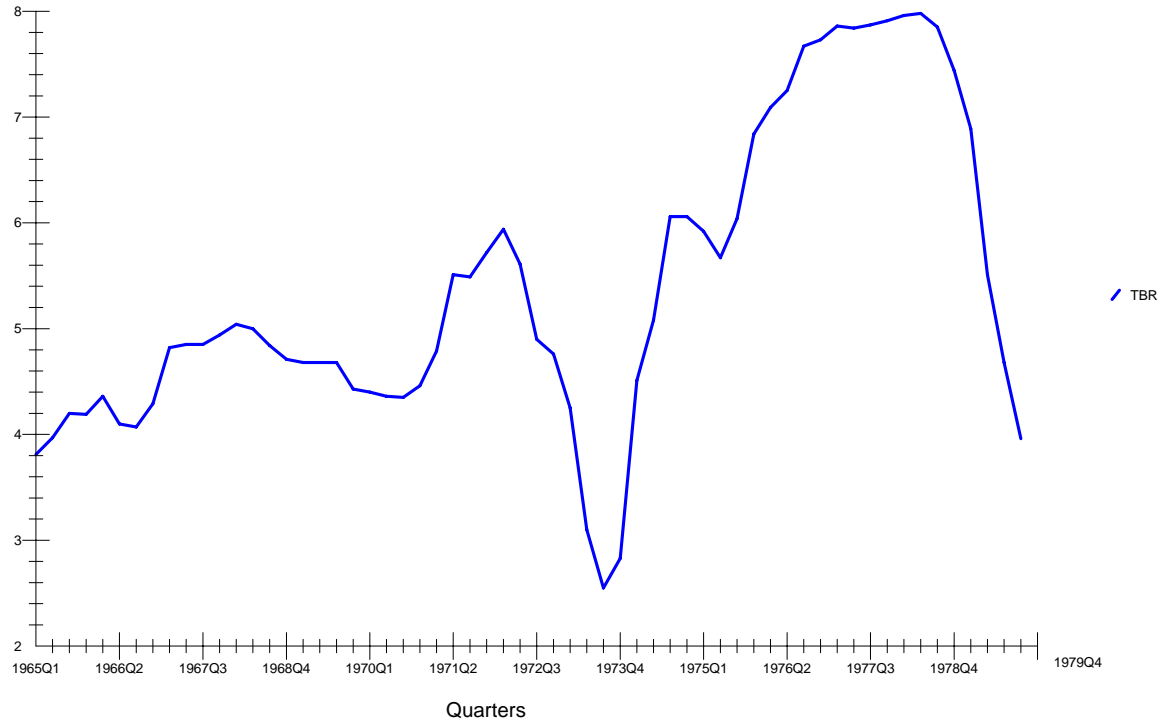
(e) Sequential test procedures can also be applied to the Perron (1994) innovational outlier and additive outlier models.

- EXAMPLE: South African Treasury Bill Rate, 1965Q1-1997Q3

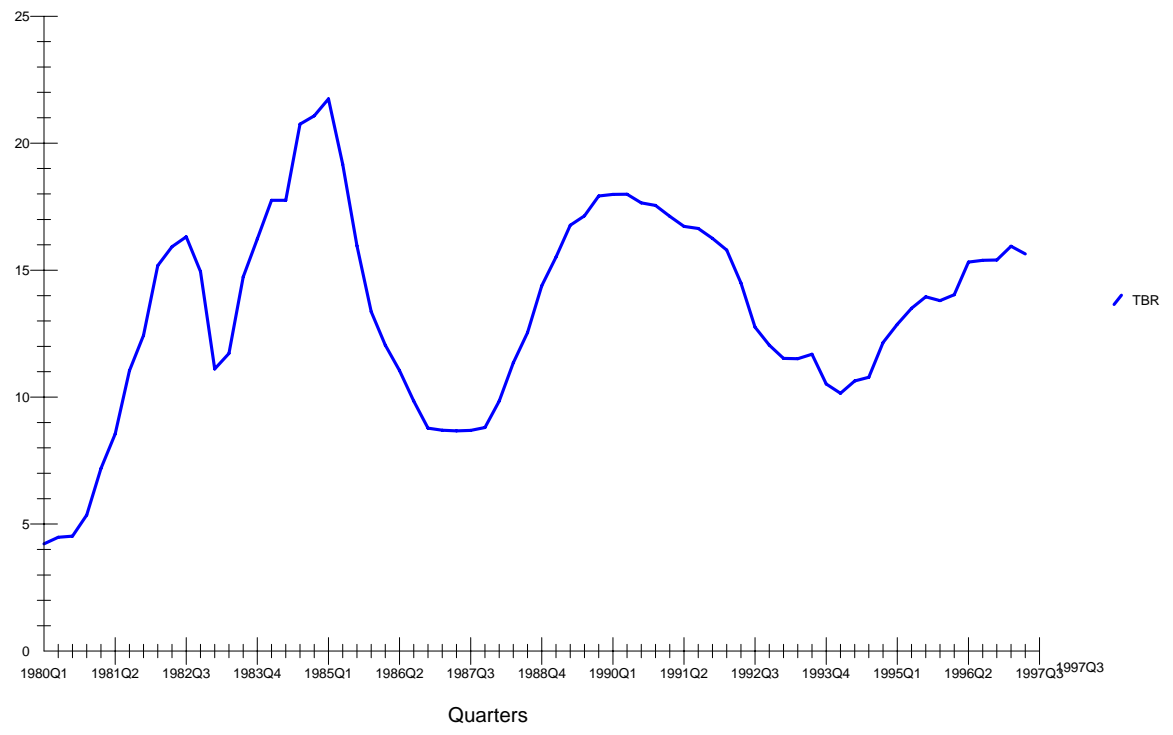


- Plausibly:
 - Structural Break, 1980Q1.
 - Monetary policy regime change: shift to market determined interest rates.
 - Stationary before & after?
 - Shift to higher mean & variance?
- Note: association with clear event

● Since:



● And:



- Also:

```

Unit root tests for variable TBR
The Dickey-Fuller regressions include an intercept and a linear trend
*****
117 observations used in the estimation of all ADF regressions.
Sample period from 1968Q2 to 1997Q2
*****
      Test Statistic      LL      AIC      SBC      HQC
DF      -1.7190      -169.7626      -172.7626      -176.9058      -174.4447
ADF(1)   -3.4902      -141.0734      -145.0734      -150.5978      -147.3162
ADF(2)   -3.6778      -140.3923      -145.3923      -152.2977      -148.1958
ADF(3)   -3.2445      -140.1532      -146.1532      -154.4397      -149.5174
ADF(4)   -3.5449      -139.0355      -146.0355      -155.7031      -149.9604
ADF(5)   -3.7210      -138.3619      -146.3619      -157.4106      -150.8475
ADF(6)   -2.9935      -137.0868      -146.0868      -158.5166      -151.1331
ADF(7)   -2.8620      -137.0868      -147.0868      -160.8977      -152.6938
ADF(8)   -3.0365      -136.5012      -147.5012      -162.6931      -153.6689
ADF(9)   -3.0680      -136.3027      -148.3027      -164.8757      -155.0311
ADF(10)  -3.6945      -133.2822      -146.2822      -164.2364      -153.5714
ADF(11)  -3.0944      -132.8765      -146.8765      -166.2117      -154.7264
ADF(12)  -2.8695      -132.8674      -147.8674      -168.5837      -156.2780
*****
95% critical value for the augmented Dickey-Fuller statistic = -3.4484
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

- \implies stationary!

- Though:

```

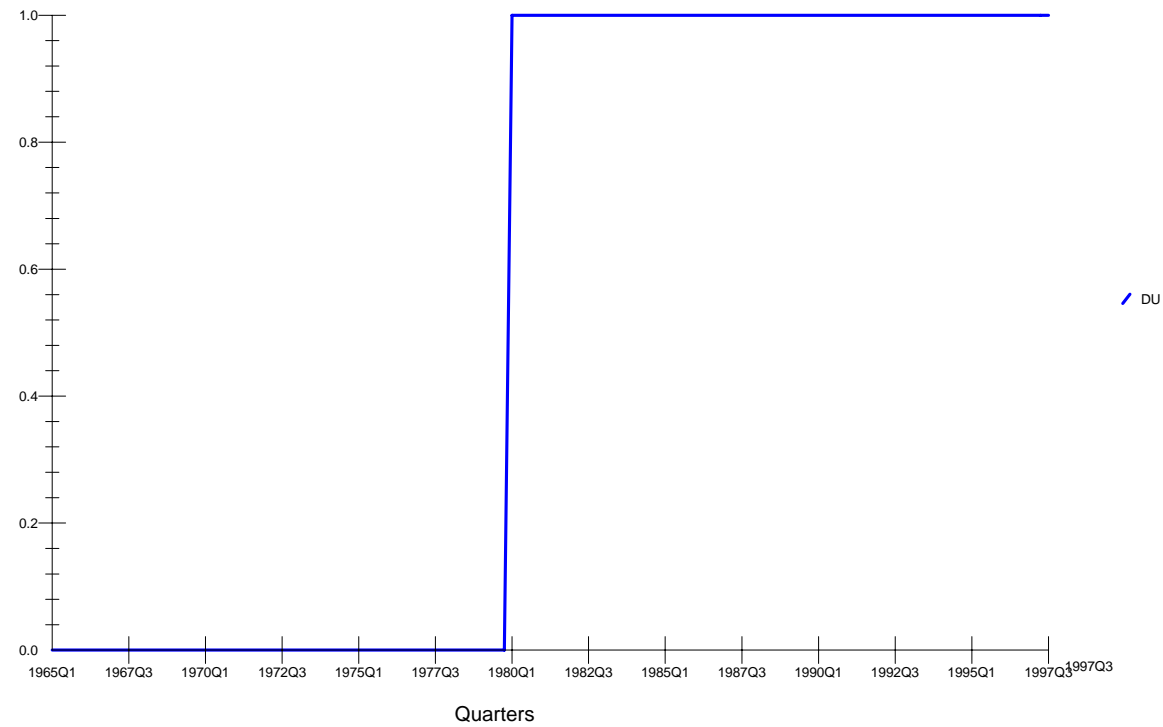
Unit root tests for variable TBR
The Dickey-Fuller regressions include an intercept but not a trend
*****
117 observations used in the estimation of all ADF regressions.
Sample period from 1968Q2 to 1997Q2
*****
      Test Statistic      LL      AIC      SBC      HQC
DF      -1.1236      -170.6212      -172.6212      -175.3834      -173.7426
ADF(1)  -2.2727      -144.4706      -147.4706      -151.6138      -149.1527
ADF(2)  -2.3118      -144.3593      -148.3593      -153.8837      -150.6022
ADF(3)  -1.9825      -143.4364      -148.4364      -155.3418      -151.2399
ADF(4)  -2.0939      -143.0978      -149.0978      -157.3843      -152.4620
ADF(5)  -2.0989      -143.0607      -150.0607      -159.7283      -153.9856
ADF(6)  -1.6741      -140.2671      -148.2671      -159.3158      -152.7527
ADF(7)  -1.5623      -140.0970      -149.0970      -161.5267      -154.1433
ADF(8)  -1.5986      -140.0019      -150.0019      -163.8128      -155.6089
ADF(9)  -1.5626      -139.9954      -150.9954      -166.1874      -157.1632
ADF(10) -1.7516      -138.8126      -150.8126      -167.3856      -157.5410
ADF(11) -1.4953      -136.8455      -149.8455      -167.7996      -157.1347
ADF(12) -1.3801      -136.3587      -150.3587      -169.6939      -158.2085
*****
95% critical value for the augmented Dickey-Fuller statistic = -2.8861
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

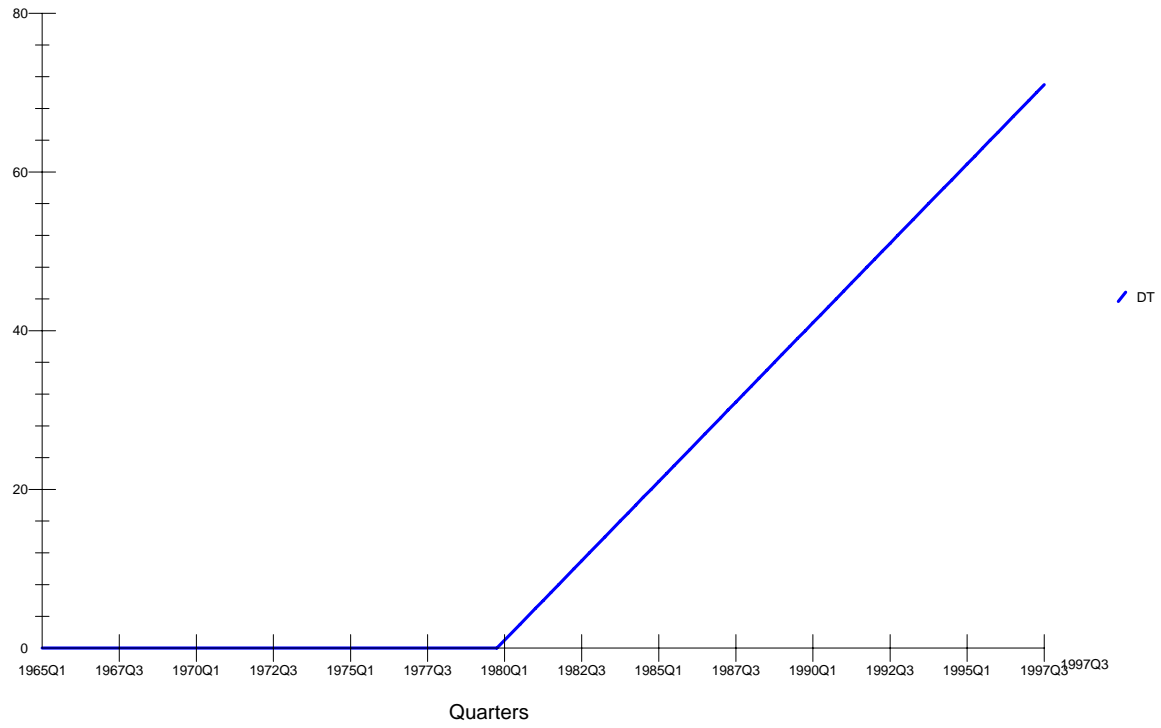
- Countervailing.

- So: Perron Innovational Outlier test: non-instantaneous effect.

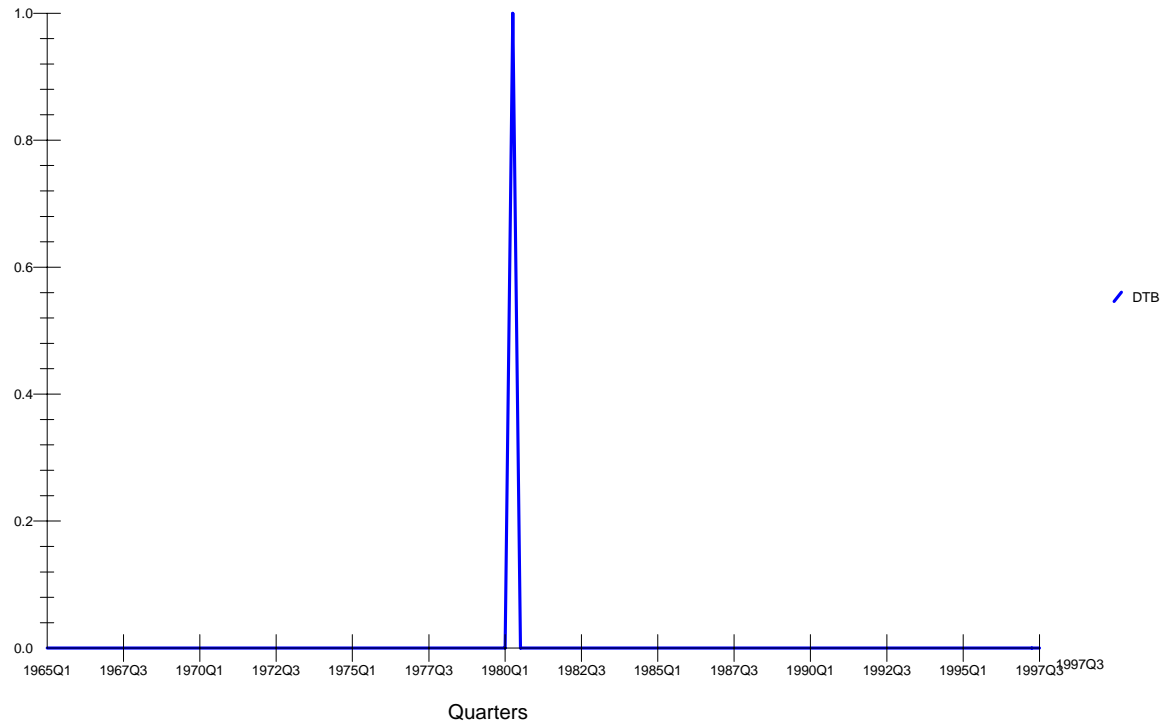
- Construct Dummies: DU:



● DT:



● DTB:



- Then estimate:

$$TBR_t = \mu + \beta t + \theta DU_t + \gamma DT_t + \delta DTB_t \\ + \alpha TBR_{t-1} + \sum \lambda_i \Delta TBR_{t-i} + e_t$$

● Gives:

```

Ordinary Least Squares Estimation
*****
Dependent variable is TBR
125 observations used for estimation from 1966Q2 to 1997Q2
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                  .52239          .25934              2.0143[.046]
T                  .0048855       .0067890           .71962[.473]
DU                 .87503         .30751              2.8455[.005]
DT                -.0028986      .0080857           -.35848[.721]
DTB               -.82478        .81821              -1.0080[.316]
TBR(-1)           .87233         .027755            31.4302[.000]
DTBR(-1)          .56373         .085792            6.5709[.000]
DTBR(-2)          .12747         .098566            1.2932[.199]
DTBR(-3)          -.10879        .099448            -1.0939[.276]
DTBR(-4)          .15658         .089820            1.7432[.084]
*****
R-Squared          .97794          R-Bar-Squared      .97621
S.E. of Regression .77431          F-stat.            F( 9, 115) 566.4095[.000]
Mean of Dependent Variable 9.9521          S.D. of Dependent Variable 5.0203
Residual Sum of Squares 68.9484          Equation Log-likelihood -140.1826
Akaike Info. Criterion -150.1826          Schwarz Bayesian Criterion -164.3242
DW-statistic       2.0774          Durbin's h-statistic -.45529[.649]
*****

```

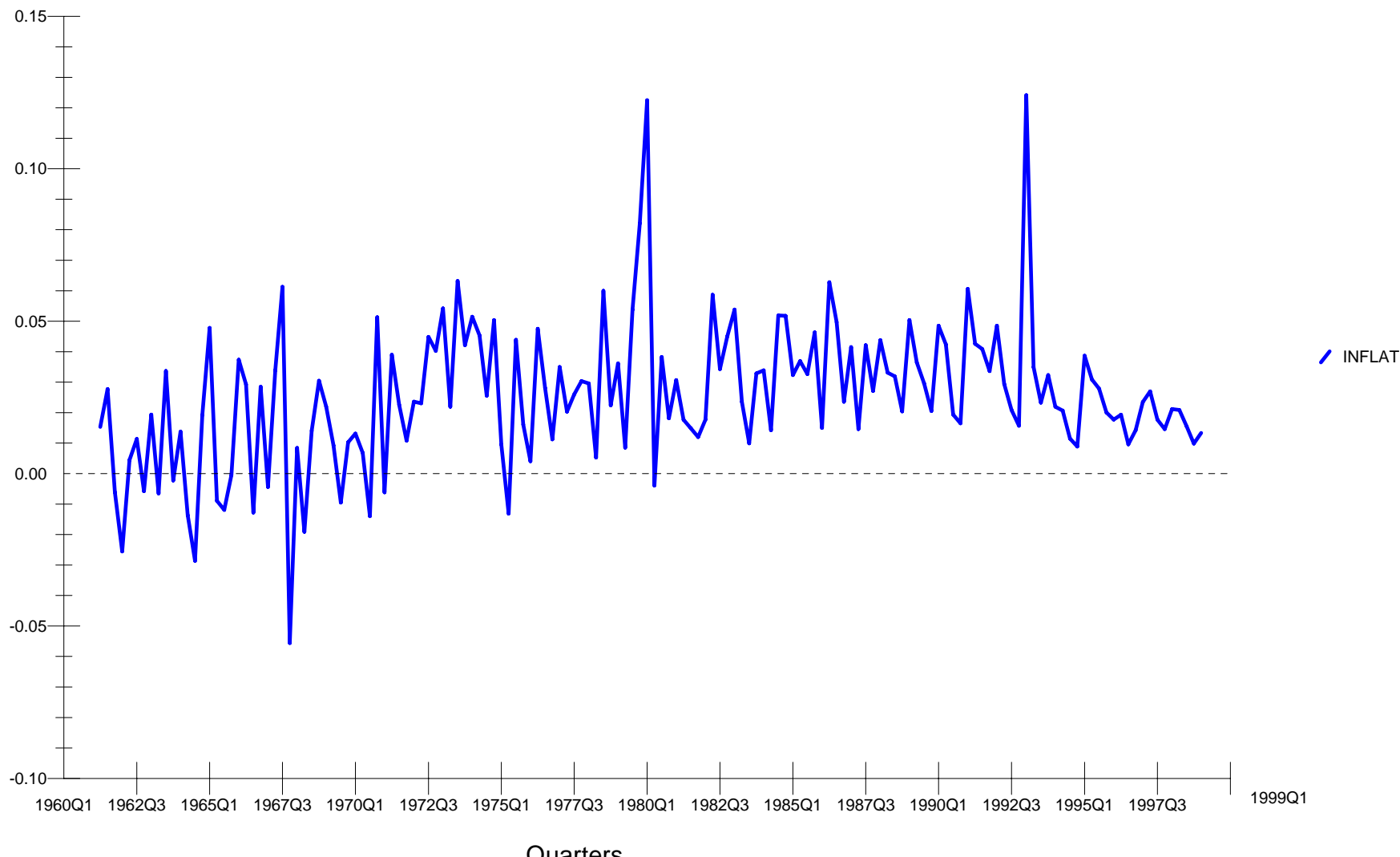
```

Diagnostic Tests
*****
* Test Statistics * LM Version * F Version *
*****
* A:Serial Correlation*CHSQ( 4)= 6.2724[.180]*F( 4, 111)= 1.4660[.217]*
*
* B:Functional Form *CHSQ( 1)= .62060[.431]*F( 1, 114)= .56881[.452]*
*
* C:Normality *CHSQ( 2)= 189.9164[.000]* Not applicable *
*
* D:Heteroscedasticity*CHSQ( 1)= 5.7721[.016]*F( 1, 123)= 5.9547[.016]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

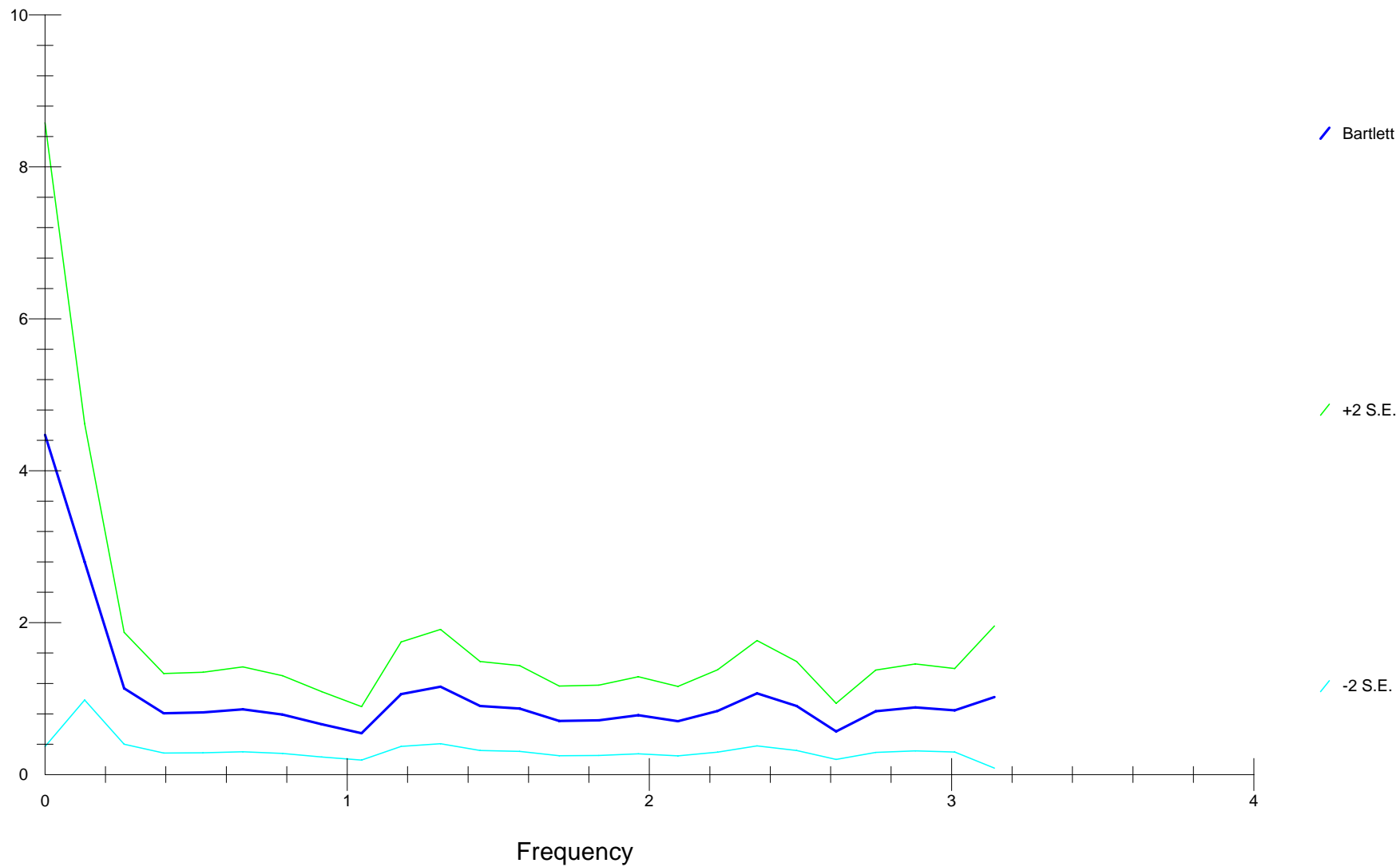
- So: $Perron = \frac{0.87233-1}{0.027755} = -4.60$
- Critical value: -5.08
- Inference: Nonstationary

- EXAMPLE: South African Inflation: CPI-based 1960Q1-1999Q1



● Given:

Standardized Spectral Density Function of INFLAT Bartlett window



- Though: ADF's ambiguous: τ_{τ} :

```

Unit root tests for variable INFLAT
The Dickey-Fuller regressions include an intercept and a linear trend
*****
139 observations used in the estimation of all ADF regressions.
Sample period from 1964Q3 to 1999Q1
*****
      Test Statistic      LL      AIC      SBC      HQC
DF      -10.6534      329.4391      326.4391      322.0374      324.6503
ADF(1)   -7.2104      329.8780      325.8780      320.0091      323.4930
ADF(2)   -5.5607      330.4227      325.4227      318.0866      322.4415
ADF(3)   -4.4221      331.2341      325.2341      316.4307      321.6566
ADF(4)   -3.5486      332.4454      325.4454      315.1748      321.2717
ADF(5)   -3.0975      332.7694      324.7694      313.0315      319.9995
ADF(6)   -3.0529      332.8348      323.8348      310.6297      318.4686
ADF(7)   -2.6553      333.2676      323.2676      308.5952      317.3051
ADF(8)   -2.3444      333.6057      322.6057      306.4661      316.0470
ADF(9)   -2.0274      334.1673      322.1673      304.5604      315.0123
ADF(10)  -1.7713      334.5761      321.5761      302.5020      313.8249
ADF(11)  -1.9270      335.0131      321.0131      300.4718      312.6656
ADF(12)  -1.8322      335.0187      320.0187      298.0101      311.0750
*****
95% critical value for the augmented Dickey-Fuller statistic = -3.4426
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

```

• And: τ_μ :

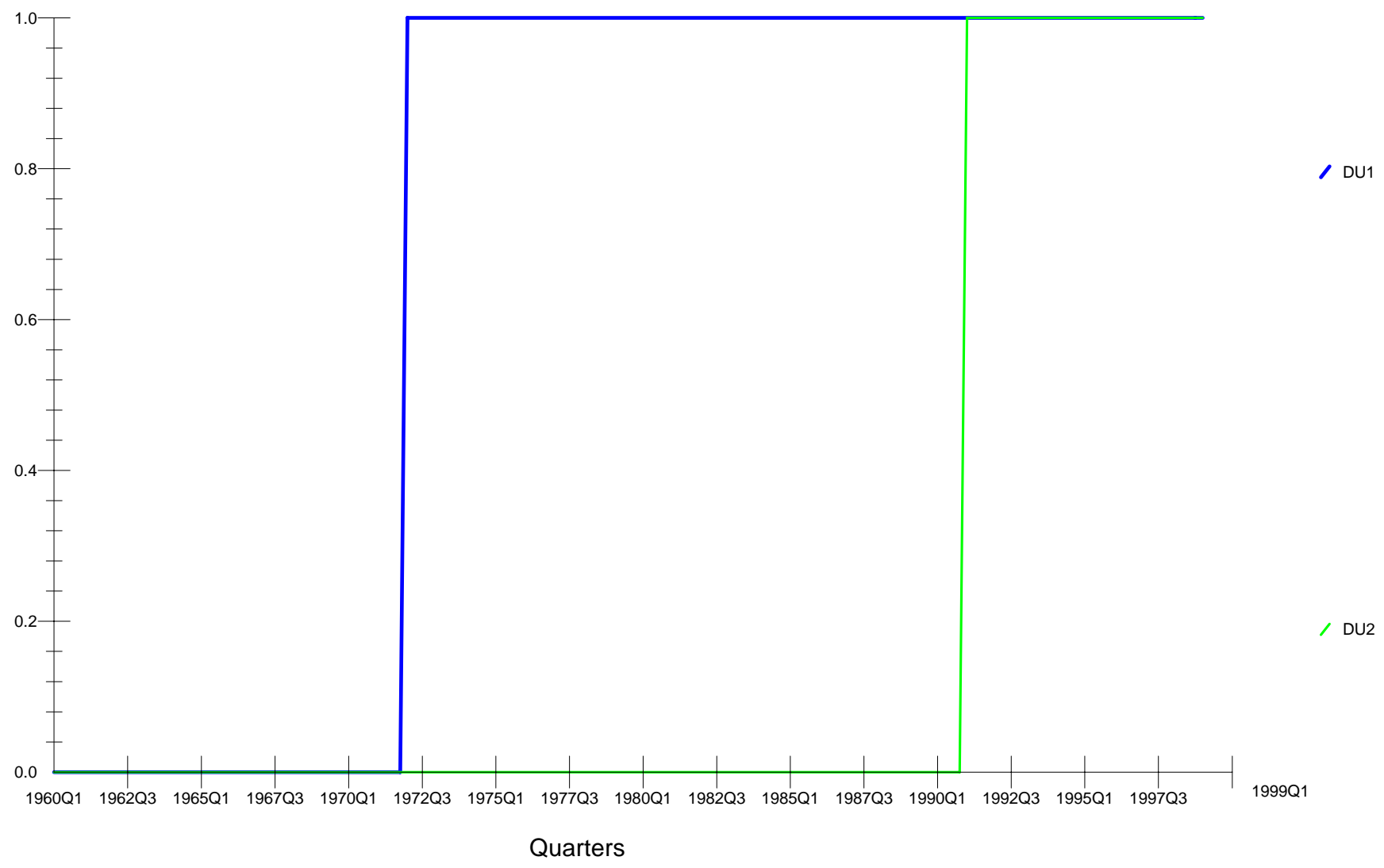
```

Unit root tests for variable INFLAT
The Dickey-Fuller regressions include an intercept but not a trend
*****
139 observations used in the estimation of all ADF regressions.
Sample period from 1964Q3 to 1999Q1
*****
      Test Statistic      LL      AIC      SBC      HQC
DF      -10.3630      327.4594      325.4594      322.5249      324.2669
ADF(1)    -6.9597      328.3362      325.3362      320.9345      323.5475
ADF(2)    -5.3660      329.2932      325.2932      319.4242      322.9082
ADF(3)    -4.2931      330.4726      325.4726      318.1364      322.4914
ADF(4)    -3.5166      332.0239      326.0239      317.2205      322.4464
ADF(5)    -3.1263      332.4859      325.4859      315.2152      321.3121
ADF(6)    -3.0763      332.5073      324.5073      312.7694      319.7374
ADF(7)    -2.7632      333.0799      324.0799      310.8748      318.7137
ADF(8)    -2.5354      333.5087      323.5087      308.8364      317.5463
ADF(9)    -2.3427      334.1459      323.1459      307.0063      316.5872
ADF(10)   -2.2118      334.5761      322.5761      304.9693      315.4212
ADF(11)   -2.2982      334.9911      321.9911      302.9170      314.2399
ADF(12)   -2.2474      335.0026      321.0026      300.4613      312.6552
*****
95% critical value for the augmented Dickey-Fuller statistic = -2.8822
LL = Maximized log-likelihood      AIC = Akaike Information Criterion
SBC = Schwarz Bayesian Criterion    HQC = Hannan-Quinn Criterion

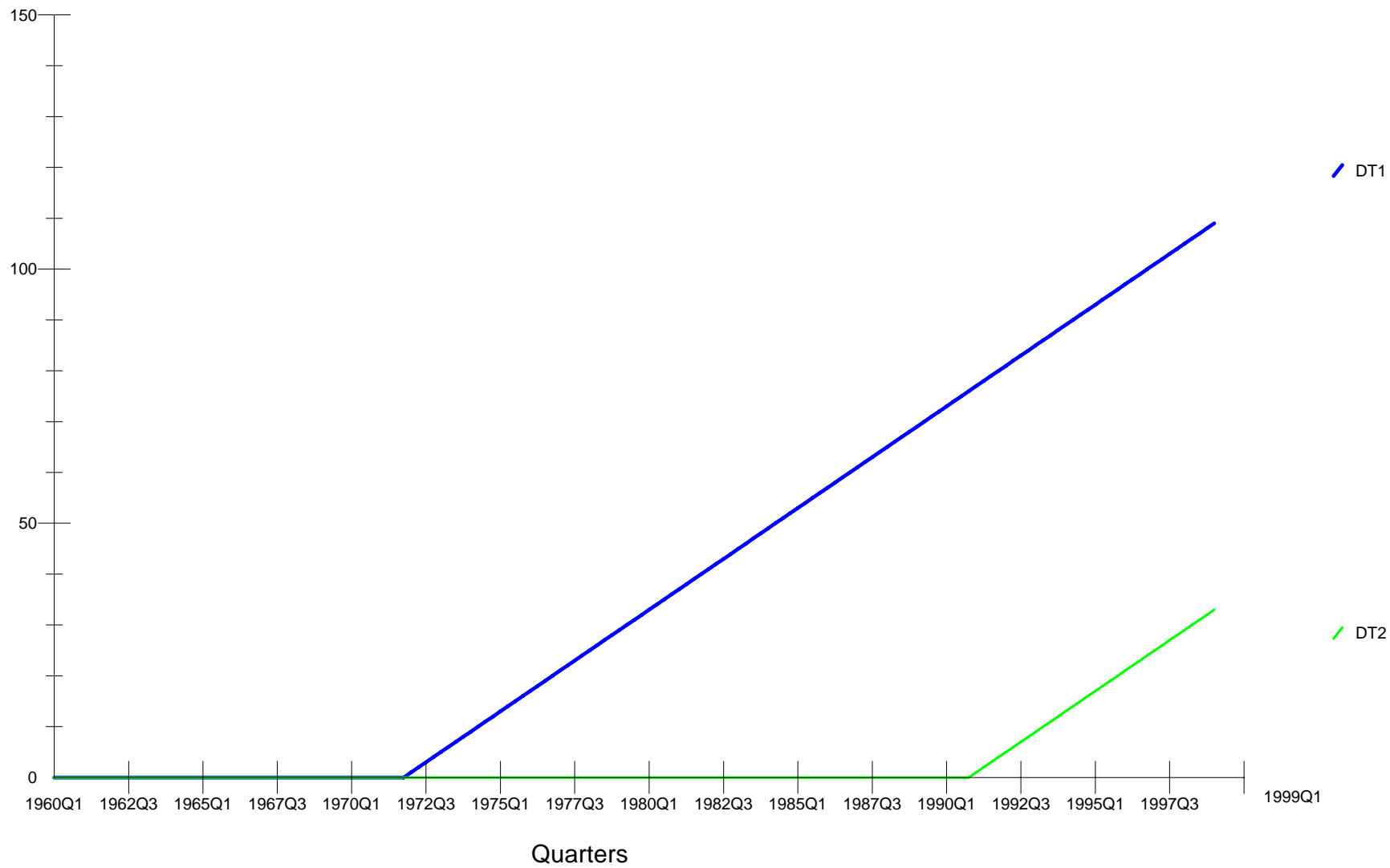
```

- Hypothesize two structural breaks:
 - 1972Q1: first oil crisis
 - 1991Q1: Stals monetary policy regime shift
- Perron Innovational Outlier test: non-instantaneous effect.
- Two sets of dummies.

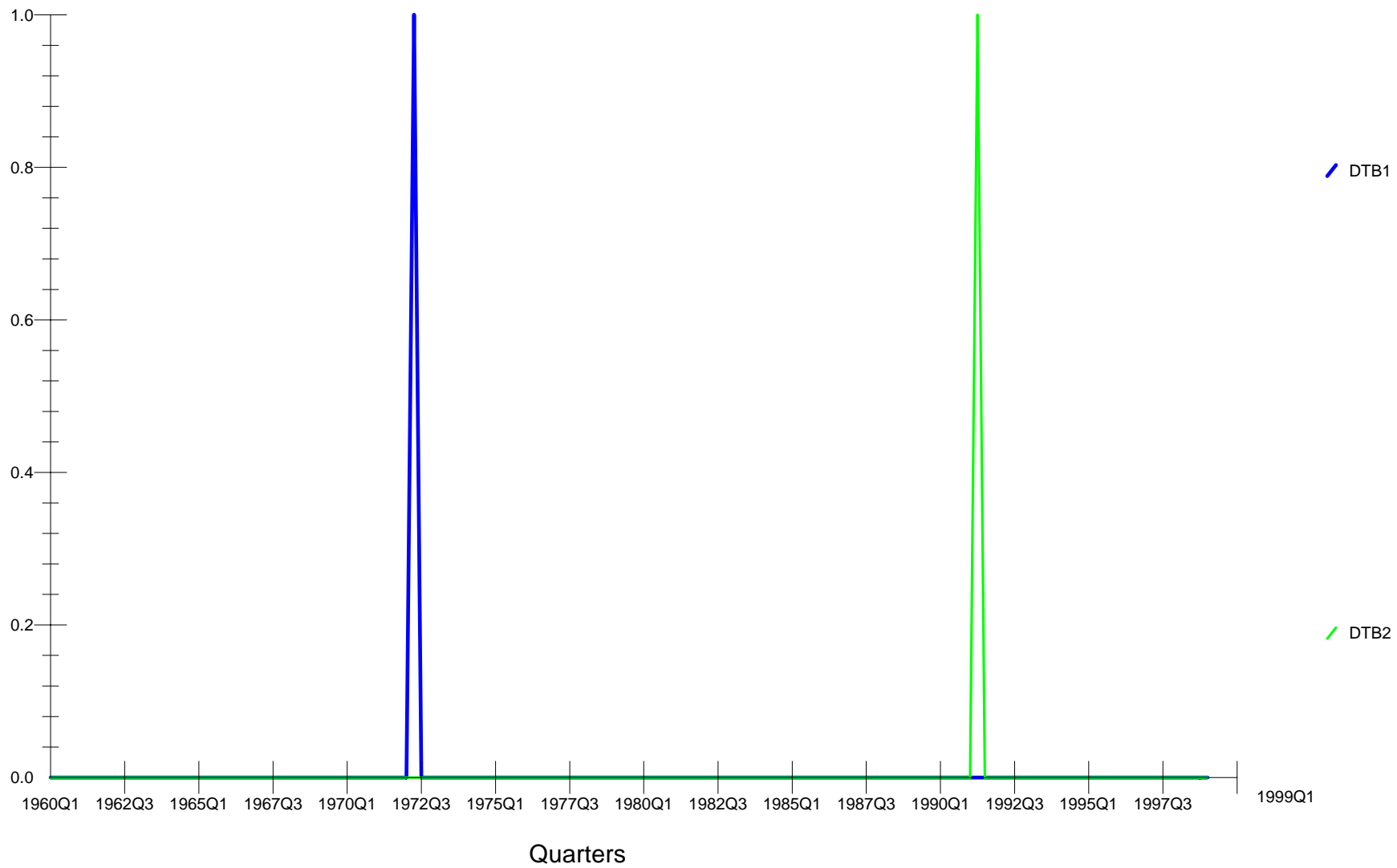
● So: DU1, DU2:



● And: DT1, DT2:



● And: DTB1, DTB2:



- Then estimate:

$$\begin{aligned} TBR_t = & \mu + \beta t + \theta_1 DU1_t + \gamma_1 DT1_t + \delta_1 DTB1_t \\ & + \theta_2 DU2_t + \gamma_2 DT2_t + \delta_2 DTB2_t \\ & + \alpha TBR_{t-1} + \sum \lambda_i \Delta TBR_{t-i} + e_t \end{aligned}$$

● To obtain:

```

                          Ordinary Least Squares Estimation
*****
Dependent variable is INFLAT
147 observations used for estimation from 1962Q3 to 1999Q1
*****
Regressor          Coefficient      Standard Error      T-Ratio[Prob]
C                  .0023699          .0098538             .24051[.810]
T                  .3585E-3          .3180E-3             1.1271[.262]
DU1                .025564           .0093560             2.7323[.007]
DT1                -.3379E-3         .3355E-3            -1.0074[.316]
DTB1               -.014883          .022071              -.67432[.501]
DU2                .015733           .0094839             1.6590[.099]
DT2                -.0014404         .4655E-3            -3.0945[.002]
DTB2               -.0036555         .022620              -.16161[.872]
INFLAT(-1)        -.37149           .22318               -1.6646[.098]
DINFLAT(-1)       .25816            .19529               1.3219[.188]
DINFLAT(-2)       .16864            .16203               1.0408[.300]
DINFLAT(-3)       .067140           .12590               .53328[.595]
DINFLAT(-4)       .029106           .085846              .33905[.735]
*****
R-Squared          .24768            R-Bar-Squared        .18031
S.E. of Regression .021133          F-stat.              F( 12, 134)          3.6764[.000]
Mean of Dependent Variable .026008          S.D. of Dependent Variable .023342
Residual Sum of Squares .059847          Equation Log-likelihood 365.1864
Akaike Info. Criterion 352.1864          Schwarz Bayesian Criterion 332.7486
DW-statistic       2.0049           Durbin's h-statistic *NONE*
*****

```

```

                          Diagnostic Tests
*****
*   Test Statistics   *           LM Version           *           F Version           *
*****
*   A:Serial Correlation*CHSQ( 4)= 1.9064[.753]*F( 4, 130)= .42702[.789]*
*   *               *           *           *           *
*   B:Functional Form *CHSQ( 1)= 1.8453[.174]*F( 1, 133)= 1.6908[.196]*
*   *               *           *           *           *
*   C:Normality      *CHSQ( 2)= 142.2319[.000]*           Not applicable           *
*   *               *           *           *           *
*   D:Heteroscedasticity*CHSQ( 1)= .019390[.889]*F( 1, 145)= .019129[.890]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values

```

- So: $Perron = \frac{-0.37149-1}{0.22318} = -6.15$
- Critical value: -5.08
- Inference: Stationary

2.6 Seasonal Unit Root Tests

- Time series data is usually presented in both seasonally adjusted, and seasonally unadjusted form.
- In general, tests for stationarity should be conducted on seasonally **unadjusted** data, since the filters which remove seasonality from the data distort the underlying properties of the data¹.
- In particular, estimates of ρ in $(A)DF$ tests tend to be biased toward 1, thus tending to find nonstationarity in the data more often than is appropriate.

- Two types of seasonal patterns:
 1. Stationary seasonal patterns show seasonal variation in a variable, but there is no persistent change in the seasonal pattern over time.
⇒ Incorporate seasonal dummies.
 2. Nonstationary seasonal patterns manifest a changing and varying seasonal pattern over time - precluding the use of seasonal dummies.
⇒ for stationarity series will require *seasonal differencing*.
⇒ seasonal processes may have more than one unit root - up to four in the case of quarterly data, for instance.

- Consider: seasonal difference operator for quarterly data is given by $\Delta_4 y_t = (1 - L^4) y_t = y_t - y_{t-4}$, where $(1 - L^4)$ can be factorized into:

$$\begin{aligned} (1 - L^4) &= (1 - L) (1 + L + L^2 + L^3) \quad (5) \\ &= (1 - L) (1 + L) (1 + L^2) \\ &= (1 - L) (1 + L) (1 + iL) (1 - iL) \end{aligned}$$

\implies 4 roots, with each root corresponding to a different cycle in the time domain:

- The first root $(1 - L)$ is the unit root we have considered thus far, at what we term the zero frequency.
- The second root $(1 + L)$ corresponds to the two-quarter, half-yearly, frequency.
- The pair of complex conjugate roots $(1 \pm iL)$ correspond to the four quarter, annual, frequency.

- Implication: test for the presence of seasonal unit roots.
- Appropriate tests are outlined by Osborne (1990).
- Note:
 - seasonal differencing involves using $(1 - L)$ to difference at the zero frequency, d , and
 - the seasonal filter $(1 + L + L^2 + L^3)$ to difference at the seasonal frequency D .
 - Thus a stochastic process y_t is said to be integrated of orders d and D , denoted $I(d, D)$, if the series becomes stationary after first period differencing d times, and seasonal differencing D times.

- Proposal, use:

$$\begin{aligned}
(1 - L) (1 - L^4) y_t &= \Delta \Delta_4 y_t \\
&= \alpha_1 D_{1t} + \alpha_2 D_{2t} + \alpha_3 D_{3t} + \alpha_4 D_{4t} \\
&\quad + \beta_1 \Delta_4 y_{t-1} \\
&\quad + \beta_2 \Delta y_{t-4} + \sum_{i=1}^{p-1} \theta_i \Delta \Delta_4 y_{t-i} + u_t \\
u_t &\sim IID(0, \sigma^2)
\end{aligned}$$

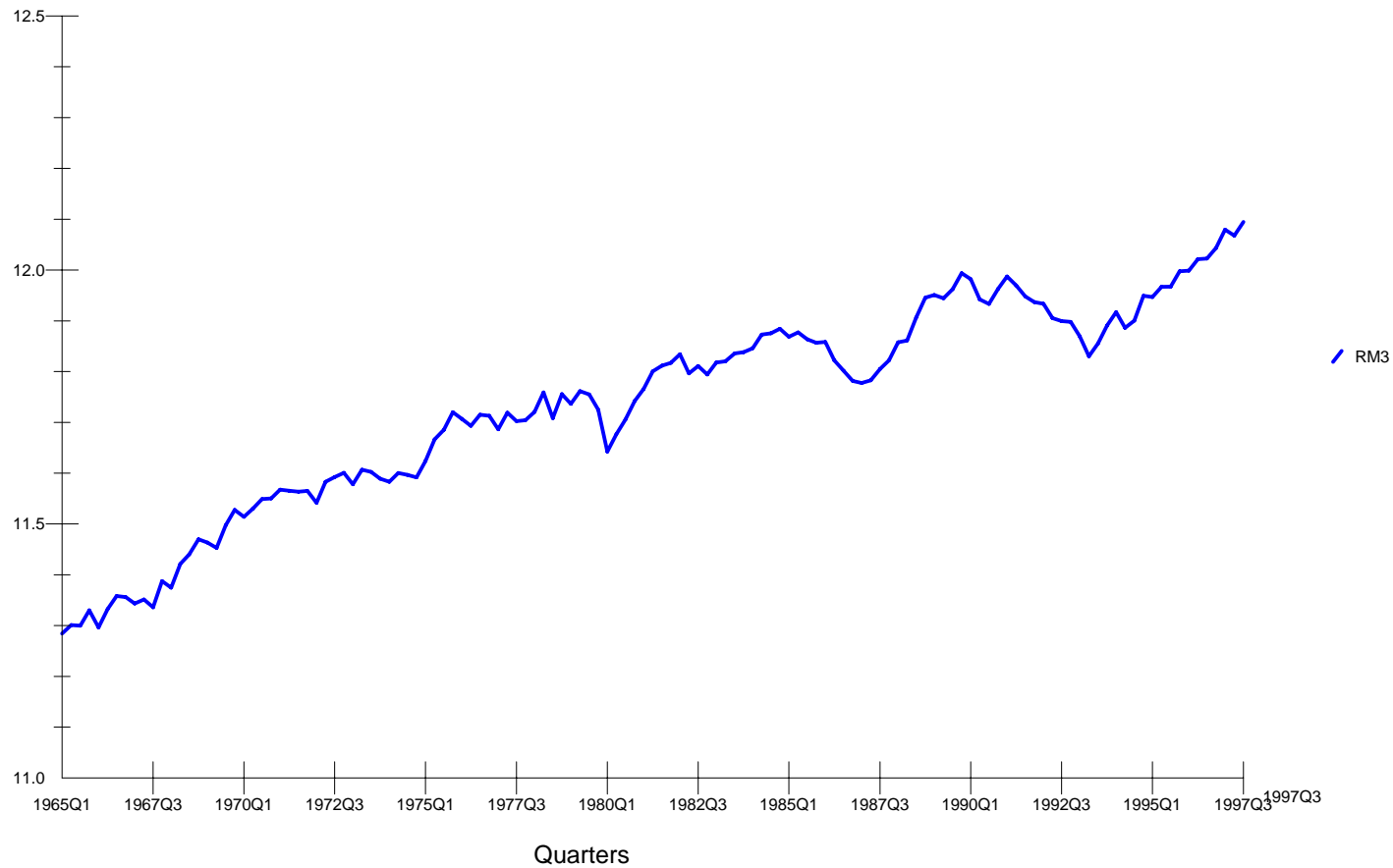
where D_{jt} is the dummy corresponding to quarter j , in order to test the null of $H_0 : I(1, 1)$.

- Where: $\beta_1 < 0$ and $\beta_2 = 0$,
 \implies acceptance of $H_a : I(0, 1)$,
- If $\beta_1 = 0$ and $\beta_2 < 0$,
 \implies acceptance of $H_a : I(1, 0)$.

- Note that:
 - Seasonal unit roots appear to be relatively rare in macroeconomic data. Osborn (1990) reports only 5 out of 30 series examined as containing a unit root.
 - Instead, most macroeconomic series appear to be $I(1)$ with superimposed deterministic seasonal patterns.

- There are good “deep” reasons for why this might be the case:
 - * If all three seasonal unit roots are present, then no two quarters will prove to be cointegrated, such that the four quarter series go their separate ways in the long run. The obvious question would be what economic series is likely to generate such a pattern.
- Even where seasonal unit roots are present, the standard *ADF* test remains valid at the zero frequency, as long as sufficient lags to take account of seasonal terms are included in the test - though some size distortion occurs.

- Example: South African Broad Real Money Balances (LOG-SCALE)



– Estimate:

$$\begin{aligned}\Delta\Delta_4RM3_t &= \alpha_1 D_{1t} + \alpha_2 D_{2t} + \alpha_3 D_{3t} + \alpha_4 D_{4t} + \beta_1 \Delta_4 RM3_{t-1} \\ &\quad + \beta_2 \Delta RM3_{t-4} + \sum_{i=1}^{p-1} \theta_i \Delta\Delta_4 RM3_{t-i} + u_t\end{aligned}$$

– Where:

$$\Delta_4 RM3 \equiv RM3_t - RM3_{t-4}$$

$$\Delta\Delta_4 RM3 \equiv D_4 RM3_t - D_4 RM3_{t-1}$$

$$\Delta RM3 \equiv RM3_t - RM3_{t-1}$$

– Where $\beta_1 < 0$ and $\beta_2 = 0$ it follows that
 $RM3 \sim I(0, 1)$,

– Where $\beta_1 = 0$ and $\beta_2 < 0$ it follows that
 $RM3 \sim I(1, 0)$.

— Gives:

```
Ordinary Least Squares Estimation
*****
Dependent variable is DD4RM3
106 observations used for estimation from 1971Q2 to 1997Q3
*****
Regressor          Coefficient          Standard Error          T-Ratio[Prob]
S1                  .0014831              .0049120                 .30193[.763]
S2                  .0055387              .0047949                 1.1551[.251]
S3                  .0061855              .0047968                 1.2895[.200]
S4                  .011754               .0049169                 2.3906[.019]
D4RM3(-1)          -.096301              .076020                  -1.2668[.208]
DRM3(-4)           -.87751               .14738                   -5.9540[.000]
DD4RM3(-1)         .12967                .084595                  1.5328[.129]
DD4RM3(-2)         .075719               .086556                  .87479[.384]
DD4RM3(-3)         .081696               .087060                  .93838[.350]
DD4RM3(-4)         .15958                .10117                   1.5774[.118]
*****
R-Squared          .45226                R-Bar-Squared          .40091
S.E. of Regression .023839              F-stat. F( 9, 96)      8.8073[.000]
Mean of Dependent Variable .1755E-3          S.D. of Dependent Variable .030800
Residual Sum of Squares .054558            Equation Log-likelihood 250.9045
Akaike Info. Criterion 240.9045          Schwarz Bayesian Criterion 227.5873
DW-statistic       1.9462
*****

Diagnostic Tests
*****
* Test Statistics * LM Version * F Version *
*****
* A:Serial Correlation*CHSQ( 4)= 4.3241[.364]*F( 4, 92)= .97814[.423]*
* B:Functional Form *CHSQ( 1)= .69475[.405]*F( 1, 95)= .62676[.431]*
* C:Normality *CHSQ( 2)= 2.6982[.259]* Not applicable *
* D:Heteroscedasticity*CHSQ( 1)= .24950[.617]*F( 1, 104)= .24537[.621]*
*****
A:Lagrange multiplier test of residual serial correlation
B:Ramsey's RESET test using the square of the fitted values
C:Based on a test of skewness and kurtosis of residuals
D:Based on the regression of squared residuals on squared fitted values
```

– Thus: $\beta_1 = \begin{matrix} -0.096301 = 0 \\ (0.076020) \end{matrix}$

– And: $\beta_2 = \begin{matrix} -0.87751 < 0 \\ (0.14738) \end{matrix}$

$\implies RM3_t \sim I(1, 0)$.

– No seasonal unit root.

2.7 Power and level of unit root tests

- Size distortion of ADF tests:
 - Over-rejection of null: too ready conclusion of stationarity
 - Distribution of ADF-tests diverges strongly from reported DF-distribution when underlying distribution contains MA element.
 - Extends to Phillips-Perron.
 - Schwert (1989), De Jong et al (1992), Agiakoglu and Newbold (1992).

- Power distortions of unit root tests:
 - Under-rejection of null: too ready conclusion of non-stationarity
 - Unit root tests low power against plausible trend stationary alternatives
 - De Jong et al (1992), Agiakoglu and Newbold (1992).

- Choosing the correct testing procedure for unit roots not always straightforward.
- *ADF* tests with different lag lengths and deterministic components can result in different outcomes with respect to testing the null of non-stationarity.
- These problems compound in small sample contexts.

- Problem:

- when we apply *ADF* tests to finite samples, it is strictly speaking inappropriate to apply critical values based on the *DF* asymptotic distribution.
- difficulty becomes most severe due to the closeness of finite sample distributions of statistics under trend-stationary and difference-stationary processes that approximate trend-stationarity.
In finite samples:
 - * Any trend-stationary process can be approximated arbitrarily well by a unit root process.
 - * Any unit root process can be approximated by a trend stationary process, especially in small samples.

- In finite samples:
 - Some unit root processes appear to behave like trend-stationary processes,
 - Some trend-stationary processes appear to behave like random walk (nonstationary) processes.
 - Unit root tests are likely to be fooled in either event: either rejecting nonstationarity where nonstationarity should be accepted, or accepting nonstationarity where nonstationarity should be rejected.

- One important consequence of this is that one's inferences concerning the stationarity properties of data should always be based on a range of evidence, of which the formal *ADF* (and other) tests are merely one component.
- Residual plots, plots of the autocorrelation function, the spectrum, etc., all form part of the evidence required to draw conclusions on the stationarity of data.